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Association between genetic obesity susceptibility and mother-reported eating behavior in children up to 5 years

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Running head

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Abbreviations

BMI-GRS: combined obesity risk-allele score

SNP: single-nucleotide polymorphisms

1 **Abstract**

2 **Background.** Many genetic polymorphisms identified by genome wide association studies
3 for adult BMI have been suggested to regulate food intake.

4 **Objective.** The objective was to study the associations between a genetic obesity risk score,
5 appetitive traits and growth of children up to age 5 years, with a longitudinal design.

6 **Methods.** In 1,142 children from the EDEN birth cohort, a combined obesity risk-allele score
7 (BMI-GRS) was related to appetitive traits (energy intake up to 12 months, a single item on
8 appetite from 4 months to 3 years, a validated appetite score at 5 years) using Poisson
9 regressions with robust standard errors. The potential mediation of appetitive traits on the
10 association between BMI-GRS and growth was assessed by the Sobel test.

11 **Results.** Children with a high BMI-GRS were more likely to have high energy intake at 1
12 year and high appetite at 2 and 5 years. High energy intake in infancy and high appetite from
13 1 year were related to higher subsequent BMI. High 2-y appetite seemed to partially mediate
14 the associations between BMI-GRS and BMI from 2 to 5 years (all $p \leq 0.05$).

15 **Conclusions.** Genetic susceptibility to childhood obesity seems to be partially explained by
16 appetitive traits in infancy, followed by an early childhood rise in BMI.

17

18 **Introduction**

19 Genome-wide-association studies (GWAs) have identified several genes associated with an
20 increased risk of obesity (1, 2). Time has now come to take advantage of these major findings
21 in order to better understand the mechanisms of lifecourse obesity development from infancy
22 to adulthood. Many birth cohorts have shown that scores of genetic susceptibility to obesity
23 are associated with early growth, as early as 2 years in some of them, but not with foetal
24 growth (3, 4, 5, 6).

25 A great part of the genes implicated by GWAs, and especially those from monogenic studies
26 on severe childhood or early onset obesity, are purportedly involved in the central regulation
27 of food intake (7). Furthermore, the genetic component of eating behavior has clearly been
28 established through heritability studies, in adults (8) and also in children (9, 10). Eating
29 behavior even in infancy may influence later adiposity development (11), which raises the
30 question whether the association between the genetic score and adiposity is mediated by
31 eating behaviour in childhood. Cross-sectional studies, conducted among 8-11 years children,
32 have already shown that increased BMI of children homozygous for the at-risk allele of a
33 FTO SNP may be partially mediated by appetitive traits (12, 13). Similarly, a cross-sectional
34 analysis of the Twins Early Development Study highlighted that satiety responsiveness
35 partially mediated the association between a polygenic risk score and adiposity in children
36 aged 8-11 years(14). However, such a cross-sectional design is insufficient to establish a
37 causal pathway. More recently, this causal pathway has been tested longitudinally within a
38 Norwegian cohort following children from the age of 4 years to the age of 8 years (15). In that
39 study, higher genetic risk for obesity was associated with appetitive traits at 6 years, but these
40 appetitive traits were not related to BMI gain up to 8 years. As a previous study has shown
41 that appetitive traits assessed at age 3 months predicted weight at 9 months, it would be
42 important to test the potential mediating effect of infant or toddler's appetitive traits on

43 subsequent growth. A prospective design would also offer the opportunity to highlight a
44 potential window of opportunity for this mediating effect.

45 In this context, we aimed at investigating within the French EDEN mother-child cohort, how
46 genetic predisposition to obesity may influence eating behavior very early life and whether
47 this relationship might explain the accelerated growth in children with higher genetic risk for
48 obesity. In the present study, we focused on three aspects of child's eating behavior that could
49 be assessed longitudinally, from 4 months to 5 years: energy intake and appetite.

50 **Material and methods**

51 *Study population*

52 The EDEN mother-child study is a prospective cohort whose aim is to assess pre- and post-
53 natal determinants of child growth, development and health. It has been described in more
54 detail elsewhere (16). In brief, 2002 pregnant women were recruited from 2003 to 2006 in
55 two French university hospitals, before 24 weeks of amenorrhea. Exclusion criteria were
56 multiple pregnancies, known diabetes prior to pregnancy, illiteracy, plan to move outside the
57 region within the next three years. The study was approved by the Ethics Committee of the
58 University Hospital of Kremlin-Bicêtre on December 12, 2002 and data files were declared to
59 the National Committee for Processed Data and Freedom. Written consent was obtained from
60 both parents.

61 The primary endpoints of the cohort were pre and post-natal growth, blood pressure, allergies,
62 infectious diseases, mental health and cognitive development. Dietary behavior was initially
63 collected as one of the main exposure factors in the cohort and was considered as a potential
64 mediating factor in the present study.

65 Data collected during pregnancy and at birth, include sociodemographic variables, maternal
66 smoking, gestational diabetes, parental anthropometric measurements, and newborn
67 characteristics (sex, gestational age, birthweight) (17). At four, eight and twelve months after
68 birth, mothers completed mailed questionnaires that provided detailed information on the
69 feeding method (exclusive breast-feeding, exclusive formula-feeding or mixed-diet) between
70 birth and four months, then at eight and twelve months as well as the date of breastfeeding
71 cessation. Breastfeeding duration was derived from these data (18). Clinical exams were also
72 conducted at birth, 1 year, 3 years and 5 years and were dedicated to collection of biological
73 sample, cognitive assessments by psychologists and measurement of anthropometric and
74 clinical parameters.

75 *Genotyping and BMI Genetic Risk Score (BMI-GRS) generation*

76 DNA was extracted from cord blood samples collected at birth. Genotypes at 16 single-
77 nucleotide polymorphisms (SNPs) were measured at the Medical Research Council
78 Epidemiology Unit, Cambridge (iPLEX platform; Sequenom), as previously described (19).
79 Among the 32 loci identified by Speliotes et al. as having genome-wide significant
80 associations with BMI in adults (2), we considered in the present study the 16 SNPs also
81 showing associations with childhood BMI either in that original report (2) or in subsequent
82 data (4); they lie in or near NRXN3 (rs10146997), SLC39A8 (rs13107325), TNNT3K
83 (rs1514175), PTBP2 (rs1555543), MC4R (rs17782313), FLJ35779 (rs2112347), NEGR1
84 (rs2568958), RPL27A (rs4929949), TMEM18 (rs6548238), RBJ/POMC (rs713586), CADM2
85 (rs7640855), TRA2B/ETV5 (rs7647305), BDNF (rs925946), TFAP2B (rs987237), FTO
86 (rs9941349), and ZNF608 (rs4836133). The summary of genotyped SNPs and allele
87 frequencies is presented in **supplementary table S1**. All variants passed genotyping quality
88 control criteria (call rate, >95%; Hardy-Weinberg equilibrium, $P > .01$).

89 Combined obesity risk-allele scores, indicating genetic susceptibility to obesity, were
90 calculated for each participant as the sum of risk alleles (0, 1 or 2 at each locus) associated
91 with higher BMI across the 16 SNP loci. The score ranged from 0 to 32. To minimize dropout
92 due to missing genotype data, infants with missing genotype data at 4 (25%) or fewer loci
93 were imputed with the mean number of susceptibility alleles in their cohort for each locus. In
94 sensitivity analyses, we also computed a weighted BMI-GRS, where each risk allele was
95 multiplied by the ‘European only sex combined’ effect estimate for the BMI-increasing allele
96 reported by Locke and colleagues (1).

97 *Child’s eating behavior*

98 Maternal perception of infant’s appetite was assessed, at ages 4, 8 and 12 months, by one
99 question, translated as follow: “Usually, you would say that your baby: 1/ is always hungry or
100 demanding to feed 2/ demands to feed the same as other babies of the same age 3/ needs to be
101 stimulated to eat”, at ages 2 and 3 years, by one question: “Usually, you would say that your
102 child: 1/ is always hungry or asking to eat 2/ has a normal appetite compared with other
103 children of the same age 3/ is not often hungry”, and at 5 years by the Low Appetite scale
104 from the Children’s Eating Difficulties Questionnaire (20). From 4 months to 4 years, a high
105 appetite was defined as the category of children “always hungry or claiming to eat” and at 5
106 years as a Low Appetite score below the 10th percentile.

107 Infant diet was assessed by food records on three non-consecutive days (two weekdays and
108 one weekend day), when the infant was not sick, at 4, 8 and 12 months. Data were checked
109 and computerized a posteriori by a dietician. Nutrient intake was then calculated based on two
110 food composition databases, one specific to ready-prepared baby foods from the French baby
111 foods industry group (SFAE 2005) and one for common foods from the French Observatory
112 of Food Nutritional Quality (CIQUAL 2006) (21). Total nutrient intake was calculated only
113 among infants who were not breastfed during the three days of the dietary records, given that

114 the amount of breast milk intake could not be measured in breastfed infants. At each age, high
115 energy intake was defined as the highest quintile of energy intake. Unfortunately, energy
116 intake was not available in 2-5-year-old children, as diet was collected from a 24-item food
117 frequency questionnaire from 2 years instead of 3-d dietary records.

118 *Child's growth*

119 At each clinical examination, child's weight and height were measured. In between, weight
120 and height data were collected from self-administered questionnaires and from measurements
121 noted in the child's health booklet by health professionals. Children had on average 22 weight
122 measurements (interquartile range 16–26) from birth to 5–6 years. Individual growth curves
123 were obtained for weight and height using the Jenss-Bayley growth curve model (22). This
124 model allows us to predict weight, length/height and then calculate BMI at any given age
125 (22). We used these model-based values of weight, length/height and BMI at 1, 2, 3, 4 and 5
126 y. At each age, the WHO growth standards were used for the calculation of weight-for-age,
127 length/height-for-age and BMI-for-age z-scores. The WHO Anthro SAS macro (WHO
128 Anthro, SAS Macro, Geneva: World Health Organization;
129 <http://www.who.int/childgrowth/software/en/>) was used for this calculation.

130 *Sample selection*

131 Of the 2002 women who were recruited, 76 women were excluded because they left the study
132 before or at the time of delivery, 24 because of miscarriages, intra-uterine death, or
133 discontinuation of pregnancy for medical reasons, and 9 because they delivered outside the
134 study hospitals. Data on birthweight were available for 1899 newborns. Among them,
135 genotyping data were available for 1324 children. Children without any data on eating
136 behavior from 4 months to 5 years (n=123) as well as those with missing data on age, sex,
137 growth, breastfeeding duration or maternal smoking (n=59) were excluded from the analyses.
138 Among the 1899 newborns with data on birth weight, the 1142 participants included in our

139 analyses were quite similar to those excluded concerning gender ($p=0.9$) and maternal pre-
140 gestational BMI ($p=0.12$) but had older mothers (29.7 vs 29.1 years at delivery, $p=0.007$) and
141 were born with higher gestational age (39.4 vs. 39.0 weeks of amenorrhea, $p<0.0001$) and
142 higher birth weight (3324 g vs. 3210 g, $p<0.0001$).

143 *Statistical analyses*

144 Student *t*-tests, χ^2 -tests and 1-way ANOVA were used to test differences between included
145 and excluded populations.

146 Associations between the obesity risk-allele score (BMI-GRS) and child's eating behavior
147 were tested by Poisson regression models with robust standard errors, adjusted for recruitment
148 center, infant's gender, and age at eating behavior assessment. Additional analyses were
149 conducted to adjust further for breastfeeding duration and maternal smoking during
150 pregnancy.

151 Associations between infant/child eating behavior and predicted child's weight-for-age,
152 length/height-for age or BMI-for-age WHO z-scores were tested by linear regression models
153 adjusted for recruitment center, infant's sex, age at eating behavior assessment. Additional
154 analyses were conducted to adjust further for breastfeeding duration, maternal smoking during
155 pregnancy and parental height.

156 Associations between BMI-GRS and predicted child's weight, length/height or BMI z-scores
157 were tested by linear regression models adjusted for recruitment center and infant's sex.

158 If an eating behaviour was associated with both the independent variable (BMI-GRS, path a,
159 Supplementary Figure S1) and the dependent variable (WHO z-score, path b), we tested for
160 the presence of mediation. Mediation is said to occur when a third variable lies on the causal
161 pathway between an exposure (in this case, the BMI-GRS) and an outcome (in this case,
162 WHO z-scores). To test for mediation, the linear regression of genetic risk score on each
163 WHO z-score was adjusted for the considered eating behaviour. The presence of mediation

164 was established using the Sobel test (23, 24) and quantified by the mediation ratio ($(\beta - \beta') / \beta$),
165 where β is the initial coefficient for BMI from the model *WHO z-score ~ BMI-GRS* (path c)
166 and β' is the coefficient for WHO z-score after the model is additionally adjusted for the
167 eating behaviour (path c'). As high score on appetitive traits were considered as binary
168 variables, we used the method adapted by Jasti et al. to calculate the Sobel test (25).
169 We conducted sensitivity analyses using a weighted genetic risk score for BMI (WtBMI-
170 GRS) instead of the crude genetic risk score. To calculate this weighted score, at each locus,
171 the number of BMI increasing variants was multiplied by the effect estimate for the BMI-
172 increasing variant from Speliotes et al (2).
173 All analyses were carried out using SAS V9.3 (SAS, Cary, NC).

174 **Results**

175 Among the 1142 mothers-child pairs with both genotyping data and eating behavior
176 assessment, mothers were aged on average 29.7 (SD=4.8) years, had a pre-pregnancy BMI of
177 23.1 (SD=4.3) kg/m² and breastfed for 3.3 (SD=3.7) months. Children carried on average 13.7
178 (SD=2.5) obesity risk alleles, with a range of 5 to 22 obesity risk alleles. At 5 years, 6.6 % of
179 children were considered overweight or obese using IOTF definition (26). Children's BMI
180 and appetitive traits are described in **Supplementary table S2**.

181 ***BMI-GRS related to eating behaviours***

182 BMI-GRS was not significantly related to high energy intake at 4 mo and 8 mo but was
183 positively related to energy intake at 1 year (in infants who were not breastfed at time of
184 energy intake assessment) (**Figure 1**). Among all children, BMI-GRS was positively related
185 to maternal perception of high appetite between 8 months and 5 years, although not
186 significant in the first year or at 3 years. Further adjustment for breastfeeding duration and
187 maternal smoking during pregnancy did not substantially modify the associations (data not

188 shown). In sensitivity analyses, when considering the weighted BMI-GRS, results were
189 similar (data not shown).

190 *Child's eating behavior and growth parameters*

191 Among children not breastfed during the 3-d dietary records, high energy intakes at 4 and 8
192 months were positively related to weight-for-age, length/ height-for-age and BMI-for-age z-
193 scores (**Table 1**) from 1 to 5 years. High energy intake at 12 months was positively related to
194 weight-for-age z-score from 3 to 5 years but was not related to length/ height-for-age or BMI-
195 for-age z-scores.

196 High appetite before 1 year was not clearly related to subsequent WHO's z-scores. High
197 appetite was not related to length/height-for-age z-score, whatever the age considered. High
198 appetite at 1 year was positively related to weight-for-age or BMI-for-age z-scores up to 5
199 years but the strength of the association decreased with age and was no more significant for
200 BMI-for-age at 5 years. High appetite at 2 or 3 years was positively related to child's
201 subsequent BMI-for-age z-score and, to a lesser extent, weight-for-age z-score. These
202 associations were mitigated by further adjustment for breastfeeding duration, maternal
203 smoking during pregnancy and parental height, the associations but remained significant,
204 except the association between high energy intake at 12 months and weight-for-age z-score
205 from 3 to 5 years (supplementary table S3)

206 *Mediation analysis*

207 The association between BMI-GRS and child's size increased with age and became
208 significant from 3 years onwards for child's weight (**Figure 2A**) and 4 years onwards for
209 child's BMI (**Figure 2B**). High appetite at 2 years was the only eating behavior related to
210 both BMI-GRS and BMI, and therefore considered as a potential mediator in the association
211 between BMI-GRS and BMI. These associations were attenuated after adjustment for high
212 appetite at 2 years (Figure 2A and 2B). The Sobel test for mediation was significant but the

213 mediation was only partial with a mediation ratio decreasing with time from 47% for 2-y BMI
214 z-score ($p=0.03$) and 35% for 3-y BMI z-score ($p=0.03$) to 28% for 4-y BMI z-score ($p=0.05$)
215 and 24% for 5-y BMI ($p=0.05$). Similar results were found with the weighted BMI-GRS (data
216 not shown).

217 **Discussion**

218 In our study, the score of genetic predisposition to develop obesity was association with a
219 higher risk of high energy intake at 1 year, and high parental-perceived appetite at 2 and 5
220 years. The associations between the BMI-GRS and BMI-for-age or, to a lesser extent, weight-
221 for-age z-scores were attenuated by further adjustment for high appetite at 2 years, whereas
222 BMI-GRS was not related to length/height-for-age z-score up to 5 years. Finally, among
223 infants not breastfed during the food records, energy intakes at 4 and 8 months were
224 positively related to child's WHO z-scores until 5 years, but were not related to the BMI-
225 GRS.

226

227 The association between genetic susceptibility to obesity and appetitive traits has been
228 previously examined in childhood but not so early in infancy. Within 8-11 years olds twins of
229 the TEDS study, Llewellyn et al. (14) showed that a BMI-GRS (based on 28 SNPs) was
230 related to both BMI and satiety sensitivity, and that the latter mediated the association with
231 BMI. However, in a study conducted by Steinsbeckk et al. within Norwegian children aged
232 from 4 to 6 years at inclusion (15), the BMI-GRS (based on 32 SNPs) was related to BMI
233 from 4 to 8 years, to BMI change from 4 to 6 years and from 6 to 8 years, but was associated
234 only with slowness in eating (negative association) at 6 years and not with other eating
235 behaviours (i.e. food responsiveness, emotional overeating, enjoyment of food or satiety
236 responsiveness). In that study, the association between the BMI-GRS and BMI was not
237 mediated by any appetitive trait assessed at age 6 years.

238 If evidence of a mediating effect of appetitive traits on the association between the BMI-GRS
239 and BMI remains scarce, many studies examined the influence of appetitive traits on BMI.
240 The associations shown in EDEN are in agreement with previous studies in spite of the
241 heterogeneity of ages and methods of investigation. In the GUSTO study (27), food
242 responsiveness at 3 months was related to higher BMI up to 15 months, whereas appetitive
243 traits assessed at 12 months were not related to BMI z-score in toddlers. Moreover, satiety
244 responsiveness in early childhood was found related to lower subsequent BMI (28, 29).
245 Similar results were found in the Gemini study (11), as all appetitive traits measured at 3
246 months of age were related to later weight, with stronger association for general appetite and
247 satiety responsiveness. In that study, the possibility of reverse causality was suggested by the
248 observation that weight at 9 months was also related to appetitive traits in toddlers, with
249 stronger association for food responsiveness, satiety responsiveness and general appetite.
250 Despite the positive associations between infant appetite and subsequent weight gain and also
251 between infant size and subsequent, the prospective influence of appetite on weight appeared
252 to be stronger than the reverse.

253 In our study, energy intake in the first months of life was related to growth throughout early
254 childhood. High energy intake could probably be considered as an indicator of rapid growth
255 during this period and rapid growth in infancy is related to higher risk of overweight later in
256 childhood in this and other cohorts (30). The influence of high energy intake on child's 5-y
257 BMI is similar to the influence of high genetic predisposition to obesity (data not shown).
258 However, notably energy intake in infancy was not related to the BMI-GRS, and might be
259 more explained by prenatal or early postnatal exposures. We have to acknowledge that energy
260 intake was assessed only among infants who were not breastfed during the three days of the
261 dietary records, given that the amount of breast milk intake could not be measured in

262 breastfed infants. Moreover, from 2 years onwards, the dietary data collection was
263 unfortunately not designed to derive energy intake.

264

265 The main strength of this study is that eating behaviour and BMI were assessed repetitively
266 and prospectively from very early in infancy (as early as birth for BMI and four months for
267 eating behaviour), which permitted us to describe in detail the longitudinal patterns of
268 associations between the BMI-GRS, eating behaviour and BMI. However, appetite was
269 assessed only by one item, whereas other assessments, maybe based on more complex scales,
270 would have allowed us to check for consistency for this trait. Moreover, we have to
271 acknowledge that eating behaviour up to 3 years was not assessed by a validated tool up to 3
272 years as such tools were not developed when the EDEN study was designed. A validated tool
273 became available and was introduced only at the 5-y follow-up. Concerning the item on
274 appetite, a similar item was used by previous studies (10, 11) and associations with this item
275 were similar to those with other appetitive traits, if not stronger. Moreover, in the EDEN
276 study, we found similar associations at 5 years, when appetite was assessed by a validated
277 scale (20). As appetitive traits were reported by parental questionnaires and not assessed by a
278 laboratory tool, this measurement could be influenced by social desirability. Finally, because
279 of the originality and the uniqueness of the collected data, our results yet rely on data from
280 only one cohort. Further studies conducted in other birth cohorts would be necessary to
281 confirm and demonstrate the generalizability of these findings. Ideally, these studies would be
282 designed with repeated use of validated tools throughout early childhood to assess eating
283 behaviour or energy intake and conducted in populations with higher prevalence of
284 overweight. As previous studies have shown that these BMI variants could to be relevant also
285 to other non-Caucasian populations (31, 32), it would be interesting to conduct similar
286 analyses in non-Caucasian populations.

287

288 **Conclusion**

289 The study provides further support from a prospective study for a partial mediation by
290 appetitive traits on the relationship between genetic susceptibility and BMI. In particular, in
291 the EDEN birth cohort, genetic susceptibility to childhood and adult obesity seems to play a
292 very precocious role by influencing appetitive traits in infants and preschool children and
293 subsequently promote faster growth. Despite the more typically recognised concerns
294 regarding infant feeding difficulties, many parents might benefit from counselling and support
295 to manage high appetitive traits in their infants and young children to prevent later obesity
296 risk.

297

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313

314 BLG and BH designed the research and wrote the manuscript. BLG, YAK, JB, AF and SC
315 analysed the data. BH and MAC oversaw the EDEN study. BH, KC, VP, KKO and MAC
316 were responsible for data collection and genotyping in EDEN. All authors reviewed drafts,
317 provided critical feedback, approved the final manuscript and were responsible for the final
318 content of the paper. BH had full access to all the data in the study and takes responsibility for
319 the integrity of the data and the accuracy of the data analysis.

320

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Figure 1. Association between obesity risk-allele score and child's eating behaviour, up to 5 years

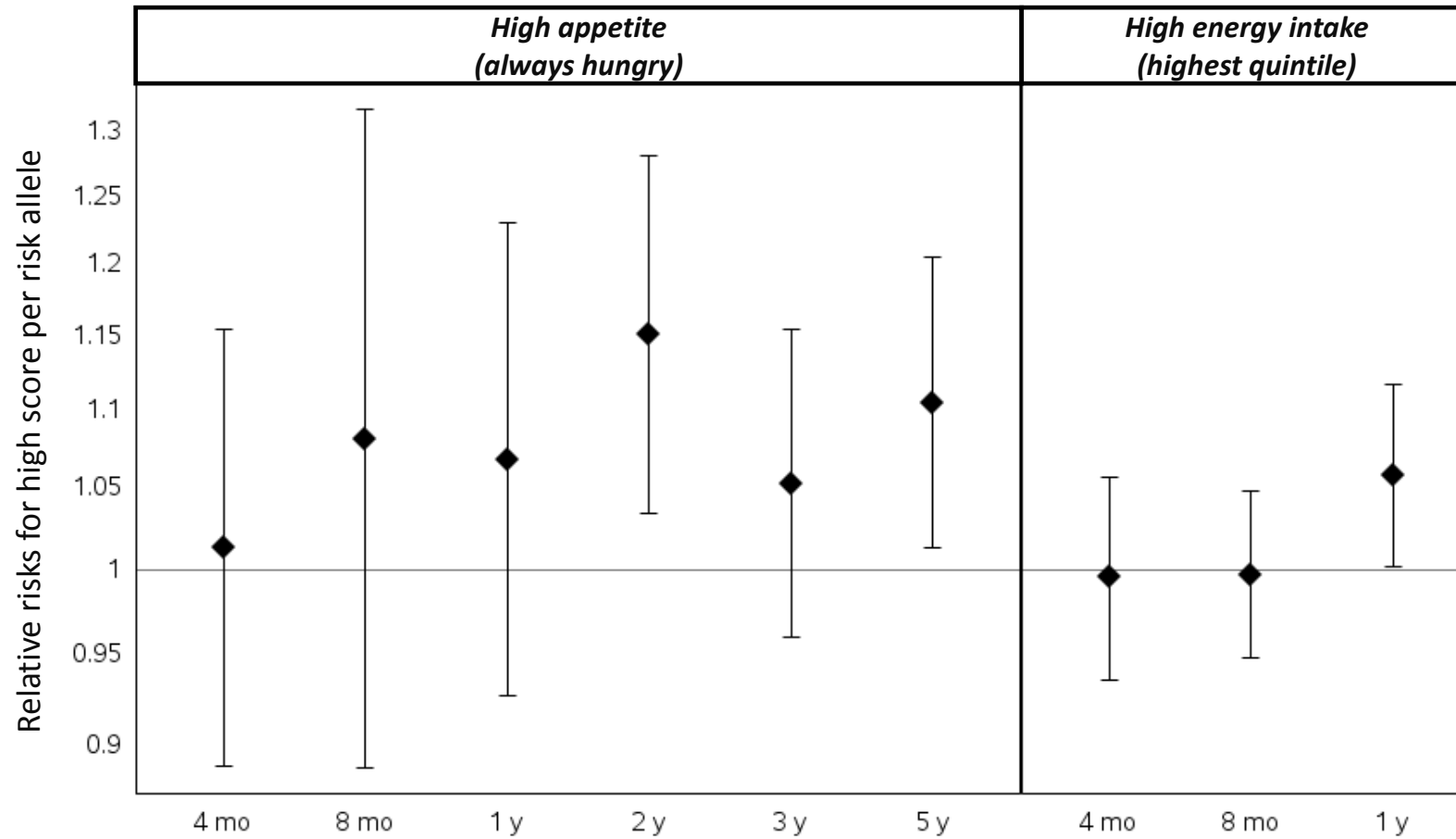


Figure legend. Results are RR [95% CI], adjusted for recruitment's center, child's age and sex (n=1081 at 4mo, 1072 at 8 mo, 1024 at 1 year, 940 at 2 years, 892 at 3 years and 790 at 5 years). Energy intake was assessed only on infants who were not breastfed during the 3-d dietary records. (679 at 4 mo, 861 at 8 mo and 822 at 12 mo)

Figure 2. Association between the obesity risk-allele score and child's WHO z-scores from 2 to 5 years

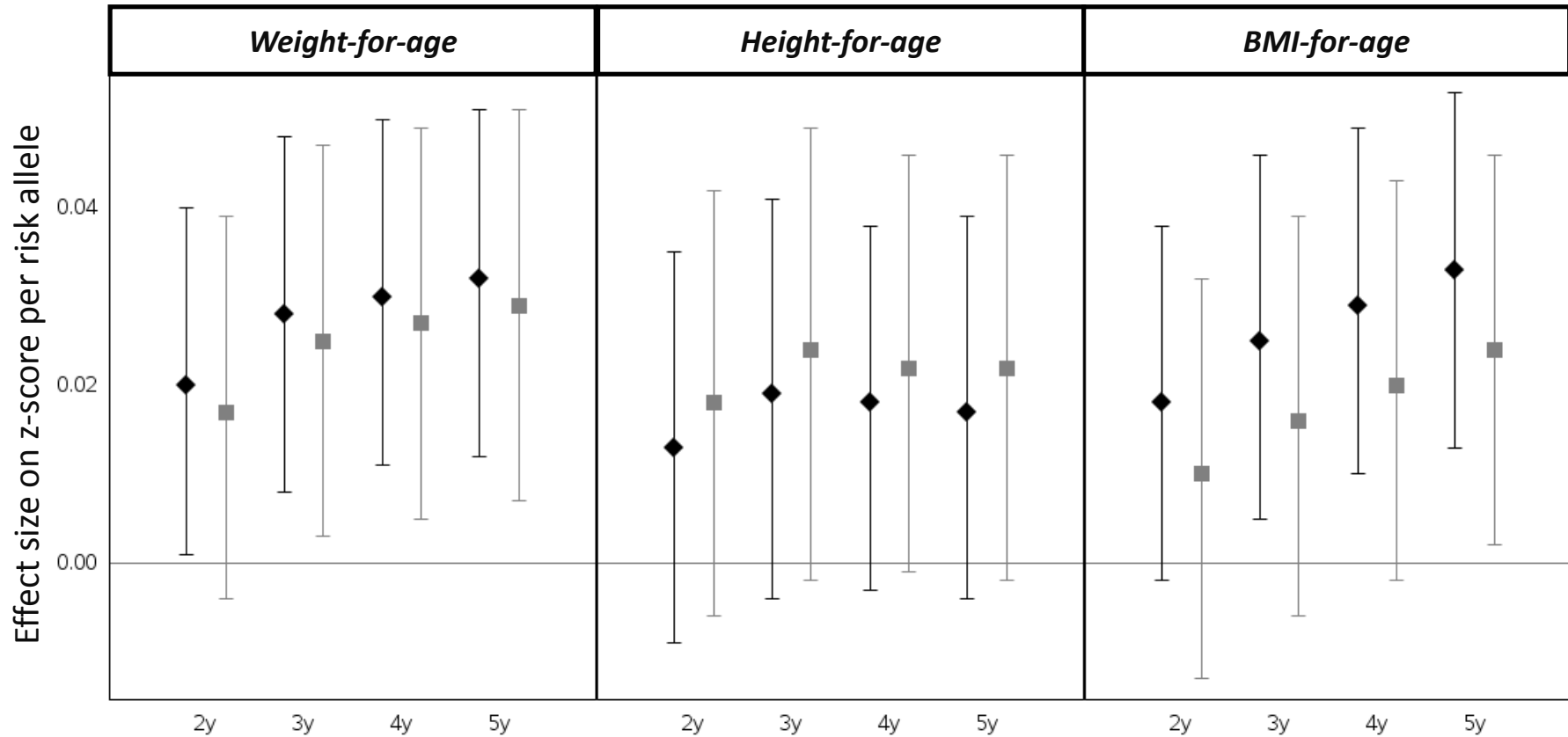


Figure legend. ◆ Results are effect size [95% CI], adjusted for recruitment center and infant's sex (n=1142). ■ Results are effect size [95% CI], adjusted for recruitment center, infant's sex, child's age and high appetite at 2 years (n=940).

SUPPORTING INFORMATION

Table S1. BMI and eating behaviour traits of children at different assessment steps

SNP	Locus	Effect allele	Other allele	SNP reported in GIANT	Evidence for association with childhood BMI	LD with GIANT SNP (r ²)	EAF in EDEN
rs10146997	NRXN3	G	A	rs10150332	Reported in Speliotes et al.	1.0	0.18
rs13107325	SLC39A8	T	C	rs13107325	Reported in Elks et al.	-	0.08
rs1514175	TNNI3K	A	G	rs1514175	Reported in Speliotes et al.	-	0.38
rs1555543	PTBP2	C	A	rs1555543	Reported in Speliotes et al.	-	0.55
rs17782313	MC4R	C	T	rs571312	Reported in den Hoed et al.	0.96	0.21
rs2112347	FLJ35779	T	G	rs2112347	Reported in Elks et al.	-	0.61
rs2568958	NEGR1	A	G	rs2815752	Reported in den Hoed et al.	0.96	0.57
rs4836133	ZNF608	A	C/G	rs4836133	Reported in Elks et al.	-	0.40
rs4929949	RPL27A	C	T	rs4929949	Reported in Speliotes et al.	-	0.45
rs6548238	TMEM18	C	T	rs2867125	Reported in den Hoed et al.	1.0	0.81
rs713586	RBJ/POMC	C	T	rs713586	Reported in Speliotes et al.	-	0.44
rs7640855	CADM2	A	G	rs13078807	Reported in Speliotes et al.	1.0	0.17
rs7647305	TRA2B	C	T	rs9816226	Reported in den Hoed et al.	0.72	0.71
rs925946	BDNF	T	G	rs10767664	Reported in den Hoed et al.	1.0	0.25
rs987237	TFAP2B	G	A	rs987237	Reported in Elks et al.	-	0.15
rs9941349	FTO	T	C	rs1558902	Reported in den Hoed et al.	0.84	0.38

Effect allele refers to the allele associated with higher BMI in the Genetic Investigation of Anthropometric Traits (GIANT) consortium (Speliotes et al., Ref. 2).

EAF: Effect allele frequency.

LD: Linkage Disequilibrium; calculated based on 1000 Genomes pilot 1 data using SNAP (<https://www.broadinstitute.org/mpg/snap/>).

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Table S2. BMI and eating behaviour traits of children at different assessment steps

	4 months	8 months	1 year	2 years	3 years	5 years
N	1081	1072	1024	940	892	790
Age (months)	4.3 (1.0)	8.3 (0.8)	12.1 (0.7)	24.4 (1.3)	37.9 (0.9)	67.7 (1.9)
Predicted BMI (kg/m ²)	16.9 (1.3)	17.1 (1.3)	16.8 (1.2)	16.2 (1.2)	15.9 (1.2)	15.4 (1.3)
High appetite (always hungry) ^a	3.7% (40)	1.9% (20)	3.1% (32)	5.1% (48)	7.2% (64)	9.0% (71)
Energy intake (kcal)						
among children not receiving breastmilk during the 3-d record	n=680 589 (104)	n=861 716 (135)	n=822 809 (135)			

Values are means (sd) or % (n).

^a at 5 years, high appetite was defined as a Low Appetite Score below the 10th percentile.

Table S3. Association between eating behaviour and child's WHO z-scores after further adjustment on parental height, maternal smoking during pregnancy and breastfeeding duration

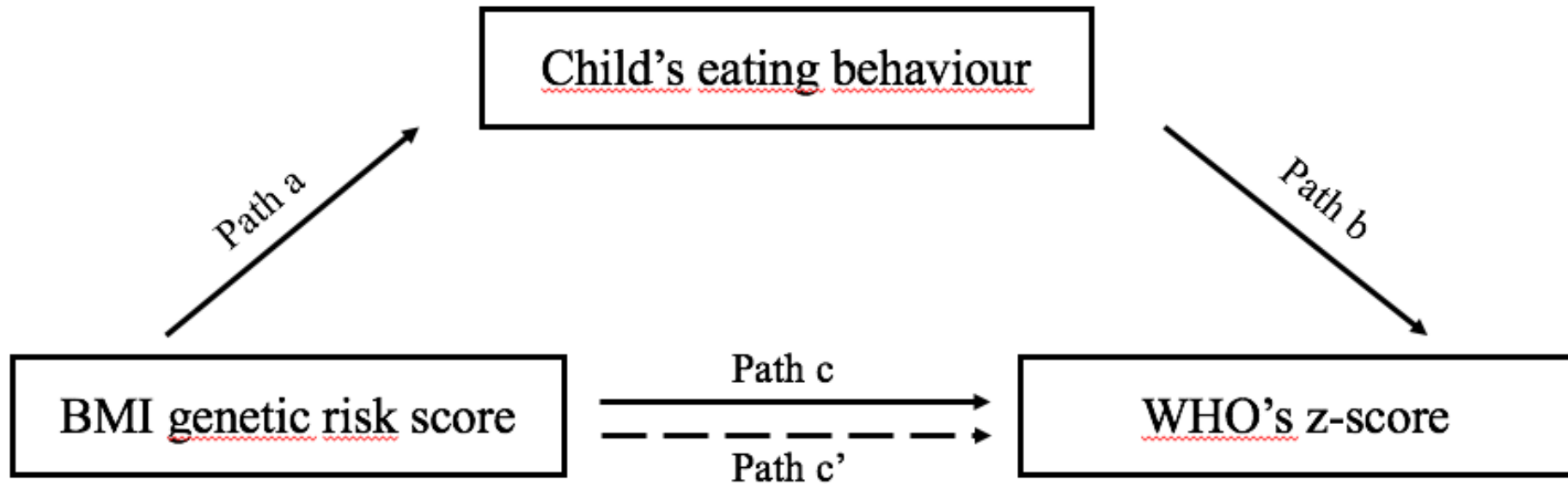
	N	Age at WHO z-score assessment				
		1 year	2 years	3 years	4 years	5 years
Outcome: Weight-for-age z-score						
High energy intake (highest quintile)						
4 months	629	0.39 [0.23 ; 0.55]	0.31 [0.15 ; 0.46]	0.25 [0.09 ; 0.41]	0.21 [0.05 ; 0.36]	0.18 [0.02 ; 0.33]
8 months	791	0.33 [0.19 ; 0.47]	0.29 [0.15 ; 0.43]	0.24 [0.10 ; 0.38]	0.19 [0.05 ; 0.33]	0.15 [0.02 ; 0.29]
12 months	767	0.04 [-0.10 ; 0.19]	0.07 [-0.07 ; 0.22]	0.09 [-0.06 ; 0.23]	0.09 [-0.05 ; 0.23]	0.09 [-0.05 ; 0.23]
High appetite (always hungry)						
4 months	993	0.32 [0.04 ; 0.59]	0.21 [-0.07 ; 0.48]	0.14 [-0.14 ; 0.42]	0.09 [-0.18 ; 0.37]	0.06 [-0.21 ; 0.34]
8 months	987	0.28 [-0.10 ; 0.66]	0.23 [-0.14 ; 0.61]	0.17 [-0.21 ; 0.55]	0.11 [-0.26 ; 0.49]	0.08 [-0.3 ; 0.45]
1 year	945	0.21 [-0.10 ; 0.51]	0.23 [-0.08 ; 0.53]	0.20 [-0.11 ; 0.51]	0.16 [-0.14 ; 0.47]	0.14 [-0.17 ; 0.44]
2 years	868		0.32 [0.08 ; 0.56]	0.28 [0.04 ; 0.53]	0.24 [0.00 ; 0.48]	0.20 [-0.04 ; 0.44]
3 years	826			0.28 [0.06 ; 0.50]	0.24 [0.02 ; 0.46]	0.21 [-0.01 ; 0.43]
5 years	738					0.41 [0.20 ; 0.61]
Outcome : Length/Height-for-age z-score						
High energy intake (highest quintile)						
4 months	629	0.25 [0.08 ; 0.42]	0.13 [-0.04 ; 0.29]	0.11 [-0.06 ; 0.28]	0.10 [-0.06 ; 0.26]	0.11 [-0.05 ; 0.27]
8 months	791	0.21 [0.06 ; 0.35]	0.20 [0.05 ; 0.34]	0.18 [0.03 ; 0.33]	0.13 [0.00 ; 0.27]	0.11 [-0.03 ; 0.25]
12 months	767	0.11 [-0.05 ; 0.26]	0.09 [-0.05 ; 0.24]	0.10 [-0.06 ; 0.25]	0.09 [-0.06 ; 0.23]	0.08 [-0.06 ; 0.23]
High appetite (always hungry)						
4 months	993	0.07 [-0.23 ; 0.37]	0.03 [-0.26 ; 0.32]	0.04 [-0.26 ; 0.34]	0.05 [-0.22 ; 0.33]	0.07 [-0.21 ; 0.35]
8 months	987	0.15 [-0.25 ; 0.55]	0.11 [-0.28 ; 0.50]	0.08 [-0.33 ; 0.48]	0.05 [-0.33 ; 0.42]	0.02 [-0.36 ; 0.40]
1 year	945	-0.27 [-0.59 ; 0.06]	-0.22 [-0.53 ; 0.09]	-0.18 [-0.50 ; 0.15]	-0.12 [-0.42 ; 0.18]	-0.07 [-0.38 ; 0.23]
2 years	868		-0.20 [-0.45 ; 0.05]	-0.20 [-0.46 ; 0.07]	-0.17 [-0.41 ; 0.07]	-0.16 [-0.41 ; 0.08]
3 years	826			0.11 [-0.12 ; 0.34]	0.09 [-0.13 ; 0.30]	0.07 [-0.15 ; 0.29]

5 years	738						0.30 [0.09 ; 0.50]
Outcome: BMI-for-age z-score							
High energy intake (highest quintile)							
4 months	629	0.35 [0.18 ; 0.52]	0.33 [0.16 ; 0.51]	0.28 [0.11 ; 0.45]	0.21 [0.04 ; 0.38]		0.16 [-0.01 ; 0.33]
8 months	791	0.29 [0.14 ; 0.44]	0.25 [0.10 ; 0.40]	0.21 [0.05 ; 0.36]	0.16 [0.02 ; 0.31]		0.14 [-0.01 ; 0.28]
12 months	767	-0.02 [-0.18 ; 0.13]	0.03 [-0.13 ; 0.18]	0.05 [-0.11 ; 0.20]	0.06 [-0.10 ; 0.21]		0.06 [-0.09 ; 0.21]
High appetite (always hungry)							
4 months	993	0.37 [0.08 ; 0.66]	0.27 [-0.03 ; 0.57]	0.17 [-0.13 ; 0.47]	0.09 [-0.20 ; 0.38]		0.03 [-0.26 ; 0.33]
8 months	987	0.27 [-0.13 ; 0.67]	0.25 [-0.16 ; 0.66]	0.19 [-0.22 ; 0.60]	0.14 [-0.26 ; 0.54]		0.11 [-0.29 ; 0.51]
1 year	945	0.49 [0.17 ; 0.81]	0.52 [0.20 ; 0.85]	0.45 [0.12 ; 0.78]	0.35 [0.02 ; 0.67]		0.27 [-0.05 ; 0.59]
2 years	868		0.64 [0.38 ; 0.90]	0.60 [0.34 ; 0.86]	0.52 [0.26 ; 0.78]		0.46 [0.20 ; 0.72]
3 years	826			0.32 [0.08 ; 0.56]	0.29 [0.05 ; 0.52]		0.27 [0.03 ; 0.50]
5 years	738						0.35 [0.13 ; 0.57]

Results are effect size [95% CI] of high level of the considered appetitive trait on WHO z-score, adjusted for recruitment centre, infant's sex, age at eating behaviour assessment, parental height, maternal smoking during pregnancy and breastfeeding duration. Significant results are highlighted in bold.

^a at 5 years, high appetite was defined as a Low Appetite Score below the 10th percentile.

Figure S1. Figure depicting the mediation analyses.



Path a shows the association between the BMI genetic risk score and eating behaviour

Path b shows the association between eating behaviour and WHO's z-score

Path c shows the relationship between the BMI genetic risk score and WHO's z-score

Path c' shows the relationship between the BMI genetic risk score and WHO's z-score, adjusted for the considered eating behaviour