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Identification of a new locus at 16q12 associated with time-to-asthma onset

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90 ABSTRACT

91 Background: Asthma is a heterogeneous disease in which age-of-onset plays an important92 role.

93 **Objective:** We sought to identify the genetic variants associated with time-to-asthma onset.

Methods: We conducted a large-scale meta-analysis of nine genome-wide association studies
of time-to-asthma onset (total of 5,462 asthmatics with a broad range of age-of-asthma onset
and 8,424 controls of European ancestry) performed using survival analysis techniques.

Results: We detected five regions associated with time-to-asthma onset at genome-wide 97 significant level ($P < 5x10^{-8}$). We evidenced a new locus in 16q12 region (near *cylindromatosis* 98 99 turban tumor syndrome gene (CYLD)) and confirmed four asthma risk regions: 2q12 (IL1RL1), 6p21 (HLA-DQA1), 9p24 (IL33) and 17q12-q21 (ZPBP2-GSDMA). Conditional 100 analyses identified two distinct signals at 9p24 (both upstream of IL33) and at 17q12-q21 101 102 (near ZPBP2 and within GSDMA). These seven distinct loci explained together 6.0% of the variance in time-to-asthma onset. In addition, we showed that genetic variants at 9p24 and 103 104 17q12-q21 were strongly associated with an earlier onset of childhood asthma (P≤0.002) 105 whereas 16q12 SNP was associated with a later asthma onset (P=0.04). A high burden of disease risk alleles at these loci was associated with earlier age-of-asthma onset (4 years 106 *versus* 9-12 years, P=10⁻⁴). 107

108 **Conclusion:** The new susceptibility region for time-to-asthma onset at 16q12 harbors variants 109 that correlate with the expression of *CYLD* and *NOD2* (*nucleotide-binding oligomerization* 110 *domain 2*), two strong candidates for asthma. This study demonstrates that incorporating the 111 variability of age-of-asthma onset in asthma modeling is a helpful approach in the search for 112 disease susceptibility genes.

113

115 Key Messages:

- 16 16q12 genetic variants are associated with time-to-asthma onset and correlate with
 CYLD and *NOD2* expressions.
- Genetic variants at 9p24 (upstream of *IL33*) and 17q12-q21 (nearby *ZPBP2* and within
 GSDMA) are associated with an earlier asthma-onset whereas variants at 16q12 are
 associated with a later asthma onset.
- Taking into account the variability of age-of-asthma onset in disease modeling can
 increase the power of identifying new genes involved in asthma physiopathology.
- 123

124 Capsule summary:

This large-scale meta-analysis of nine genome-wide association studies identified 16q12
genetic variants associated with time-to-asthma onset that correlate with *CYLD* and *NOD2*expressions, two strong candidate genes implicated in inflammation.

128

Keywords: Asthma, age-of-onset, genetics, genome-wide association study, survival analysis,
conditional analysis, *CYLD*, *NOD2*

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132 Abbreviations:

- 133 *CYLD*: Cylindromatosis (turban tumor syndrome)
- 134 eQTL: Expression quantitative trait locus
- 135 *GSDMA*: Gasdermin A
- 136 GWAS: Genome-wide association study
- 137 *HLA-DQA1*: Major histocompatibility complex, class II, DQ alpha 1
- 138 HR: Hazard ratio
- 139 *IL1RL1*: Interleukin 1 receptor-like 1

- 140 *IL33*: Interleukin 33
- 141 LCL: Lymphoblastoid cell line
- 142 LD: Linkage Disequilibrium
- 143 *NFkB1*: Nuclear factor of kappa light polypeptide gene enhancer in B-cells 1
- 144 *NOD2*: Nucleotide-binding oligomerization domain containing 2
- 145 QC: Quality control
- 146 SNP: Single nucleotide polymorphism
- 147 TAO: time-to-asthma onset
- 148 *ZPBP2*: Zona pellucida binding protein 2
- 149

150 INTRODUCTION

The prevalence of asthma has dramatically increased over the past decades in high-income countries, affecting 5–16% of people worldwide.¹ It is the most common chronic disease among children and a decrease in the age-of-asthma onset has been recently documented.²

Asthma is a complex and heterogeneous disease with variable clinical expression over the life 154 span.¹ It is now well recognized that asthma is not a single disease but rather a collection of 155 different phenotypes which may represent different manifestations of a common underlying 156 pathological process or may be separate disease entities.³ One of the simplest characteristic 157 that can be used to differentiate disease phenotypes is the age at onset.^{4,5} Indeed, asthma 158 displays different characteristics according to the lifetime period at which it occurs.⁶ Early age 159 of onset is more frequently associated with a family history of asthma, allergy sensitization 160 and clinical response to triggers, whereas late-onset disease is associated with eosinophilic 161 inflammation and obesity, is more common in women, and is generally less allergic.³ 162

The risk of developing asthma has a strong genetic component, with estimated heritability 163 ranging from 35 to 95%.⁷ Genome-wide association studies (GWASs) have been successful in 164 identifying more than 20 loci associated with asthma.⁸ However, the genetic factors identified 165 to date account only for a small part of the genetic component of the disease.¹ This hidden 166 heritability might be linked to the phenotypic heterogeneity of asthma.⁹ The vast majority of 167 168 GWASs conducted until now have analyzed asthma as a binary phenotype. A few genetic studies have considered more specific definition of asthma incorporating the age of disease 169 onset. A genome-wide linkage screen conducted for time-to-asthma onset in French families 170 revealed two regions, 1p31 and 5q13, potentially linked to this phenotype.¹⁰ A single GWAS 171 has been performed on age-of-asthma onset in asthmatic children and led to the identification 172 of two loci, not found by the previous asthma GWASs; these loci on chromosomes 3p26 and 173 11q24 were associated with an earlier onset of childhood asthma.¹¹ Moreover, the effect of 174

175 17q12-q21 genetic variants, identified by the first GWAS of asthma,¹² was found to be
176 restricted to early-onset asthma.^{13,14}

177 Instead of stratifying the data according to age-of-onset of disease using an arbitrary 178 threshold, one can integrate the age-of-onset in modeling asthma risk, by using survival 179 analysis methodology applied to both asthmatic and non-asthmatic subjects. The goal of the 180 present study was to identify the genetic determinants underlying time-to-asthma onset in a 181 large meta-analysis of 5,462 asthmatics and 8,424 controls from nine independent European-182 ancestry populations.

184 METHODS

185 **Populations**

We studied 13,886 individuals of European ancestry from nine independent studies (one birth cohort, five population-based and three family studies) which were part of the GABRIEL European consortium on the genetics of asthma.¹⁴ A brief description of these studies with appropriate references is provided in the online repository and in Table E1. All of these studies had the age of asthma onset and imputed genetic data available.

191 For all studies, ethical approval was obtained from the appropriate institutional ethic192 committees and all individuals or child's legal guardians provided written informed consent.

193

194 Time-to-asthma onset definition

195 Definition of asthma was based on report of doctor's diagnosis and/or on standardized 196 questionnaires (see online repository). To model time-to-asthma onset (TAO), we used age of 197 onset or age at first wheeze for individuals who developed asthma, while in individuals who 198 were free of disease upon examinations, we used age at last examination.

199

200 Genotyping

Genotyping, single nucleotide polymorphism (SNP) imputation process and quality control 201 (QC) criteria (for individuals and SNPs) for each study are described in Table E1 in the online 202 repository. All datasets were genotyped at Centre National de Génotypage (CNG, Evry, 203 France), as part of the European GABRIEL asthma consortium.¹⁴ Quality control and 204 imputations were performed independently for each study. Genome-wide imputations were 205 conducted using MACH 1.0 software,¹⁵ with reference haplotype panels from HapMap 2. 206 SNPs with imputation quality score (Rsq) ≥ 0.5 and minor allele frequency $\ge 1\%$ were kept 207 208 for analysis. Then, to further investigate the regions associated with time-to-asthma onset at the genome-wide significant level, we used imputed data from the 1000 Genomes Project andapplied the same SNP QC criteria.

211

212 Statistical analysis and strategy of analysis

After the study-specific QC, a total of 13,886 individuals from the nine cohorts were included 213 in the present study. In each dataset, association between time-to-asthma onset and individual 214 SNPs was investigated under an additive genetic model using a Cox proportional-hazards 215 regression model adjusted for sex and the first four principal components to account for 216 population structure. A robust sandwich estimation of the variance¹⁶ was used in family data 217 to take into account familial dependencies. Moreover, due to the complex sampling design of 218 the GABRIELA study, survey regression techniques were used for this study to estimate 219 220 robust standard errors ('svy' command in Stata). Proportional hazard assumptions for the main SNP effect were tested and never rejected. Genome-wide association studies of time-to-221 222 asthma onset were first conducted in each of the nine datasets separately, and then combined through meta-analysis in order to increase power and to obtain more robust findings. Meta-223 analyzed hazard-ratios (HRs) and 95% confidence intervals (CIs) were calculated using a 224 225 fixed-effect (inverse variance) model. To assess heterogeneity of SNP effect across studies, the Cochran's Q statistic was calculated. If heterogeneity was evidenced, a random-effect 226 model was fitted. All analyses were performed using Stata® version 13.1 (STATA Corp., 227 College Station, Texas, USA). After the meta-analysis, we only kept SNPs for which at least 228 66% of the studies (≥6 studies among the 9 studies) participated to the summary statistics in 229 the meta-analysis, to reduce the rate of false-positive findings. The meta-analysis results were 230 obtained for a total of 2,387,926 SNPs. We used the classical threshold of P-value $\leq 5 \times 10^{-8}$ to 231 declare a meta-analyzed SNP effect as genome-wide significant. 232

234 Conditional analysis to uncover distinct signals at TAO-associated loci

To identify distinct TAO-associated SNPs in each region harboring genome-wide significant 235 signals, we re-analyzed separately these regions in each of the nine studies. For that purpose, 236 we added the region top SNP into the primary Cox model as a covariate and tested the effect 237 of each other SNP of that region. Then, the results were meta-analyzed using the same 238 strategy as the primary GWASs. If a secondary signal was detected in a region, a second run 239 of conditional analyses was performed to check for a third distinct signal in that region. 240 Length of explored regions was based on regional association plots and ranged from 200 kb to 241 500 kb depending on recombination hotspots. 242

243

244 *eQTL analysis and functional annotations*

We queried whether significant SNPs (or their proxies) associated with time-to-asthma onset at $P \le 5 \times 10^{-8}$ and potentially secondary signals from conditional analysis were expression quantitative trait loci (eQTLs). We used existing eQTL databases in multiple tissues (especially blood and lung) for populations of European ancestry (see online repository).¹⁷⁻²³ Functional annotations of significant SNPs (or their proxies) were obtained using ENcyclopedia Of DNA Elements (ENCODE) data²⁴ provided by the HaploReg tool.²⁵

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252 Relationship of TAO-associated loci with age-of-asthma onset

In a first step, we investigated in asthmatics whether each of the SNPs associated with TAO were also associated with age-of-asthma onset using a non-parametric rank test followed by a non-parametric equality-of-medians test. In a second step, we assessed the cumulative effect of risk alleles of SNPs found associated with the age-of-asthma onset at step 1. For that purpose, we used either the number of risk alleles or the quintiles of a polygenic score distribution. The polygenic risk score is the weighted sum of the number of age of asthma onset associated alleles with weight being the log of adjusted hazard ratio estimated in asthmatics only. The associations were tested in eight studies for which we had access to raw data (all datasets except ALSPAC) using a cox proportional hazard model adjusted on sex and principal components.

264 **RESULTS**

265 Description of populations

A total of 13,886 subjects were included in the present study (5,462 asthmatics and 8,424 non-asthmatics). Asthmatics had a mean age-of-asthma onset of 12.5 years (ranging from 0.5 to 75 years, Figure E1), a mean age of 26.8 years at examination (mean per study ranging from 9.1 to 51.3 years) and 52.6% were males. Non-asthmatics had a mean age of 32.4 years at examination (mean per study ranging from 8.9 to 55.8 years) and 49% were males (Table E1).

272

273 Genetic variants associated with time-to-asthma onset

274 The Manhattan and the Q-Q plots of the meta-analysis of time-to-asthma onset GWAS results are shown in Figure 1 and Figure E2 in the online repository respectively. A total of 155 275 276 SNPs were associated with time-to-asthma onset at the genome-wide significance level of $P < 5x10^{-8}$. These SNPs clustered into five distinct chromosomal regions (Table I), that 277 included a new risk locus on 16q12 (nearby CYLD (cylindromatosis turban tumor syndrome), 278 1 SNP) and four established risk loci for asthma: 2q12 (IL1RL1-IL18R1, 7 SNPs), 6p21 279 (nearby HLA-DOA1, 1 SNP), 9p24 (flanking IL33, 25 SNPs) and 17q12-q21 (121 SNPs 280 281 spanning 389 kb, and with the main signal located near ZPBP2). The regional association plots for these genome-wide associated loci are shown in Figures 2 and E3, and the forest 282 plots for the top signal in each region are shown in Figure E4. Three additional loci were 283 associated with time-to-asthma onset at a suggestive significance threshold $(5 \times 10^{-8} < P < 10^{-6};$ 284 Table I): MAP4K4 (2q11-q12), RORA (15q22) and IL4R (16p12-p11). 285

To determine whether any of the five TAO loci harbored additional association signals, we performed conditional association analysis in each region. For this analysis, a threshold of $P \le 2.1 \times 10^{-5}$ was used to declare significance, corresponding to a Bonferroni threshold for

2,382 independent tests. These analyses evidenced two secondary signals (Table II and Figure 289 E5 in the online repository): 1) rs413382 in 9p24 region at 73 kb of *IL33* (P= 9.7×10^{-6} after 290 conditioning on top SNP and $P=5.9 \times 10^{-8}$ in the primary meta-analysis) and 2) rs3859192 in 291 17q12-q21 region within GSDMA (P= 4.0×10^{-6} after conditioning on top SNP and P= 1.5×10^{-13} 292 in the primary meta-analysis). In contrast, at 2q12, 6p21 and 16q12 regions, the inclusion of 293 the most significant time-to-asthma onset GWAS SNP as a covariate in association analysis 294 resulted in nearly complete reduction of the association signal in these regions, suggesting 295 that there was no evidence for a second distinct genetic factor in these regions. 296

To obtain a denser map of the new TAO 16q12 locus, we repeated association analyses using 1000 Genomes Project imputed SNPs. These analyses strengthened our original finding with additional signals $(3.8 \times 10^{-8} \le P \le 2.6 \times 10^{-7})$ located in an intergenic region encompassing the lead SNP rs1861760 (Table E2 and Figure E6). These SNPs were in moderate to high LD with rs1861760 ($0.71 \le r^2 \le 0.81$) and thus did not represent independent signals from that top hit. Similar analyses conducted in the four other TAO-associated regions also supported our original findings and did not evidence any additional independent signal in these regions.

Overall, the seven distinct SNPs (five top SNPs and two secondary SNPs) associated with TAO showed low heterogeneity between studies (P>0.11) and explained together 6.0% of the variance in time-to-asthma onset.

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308 Functional annotations and effect on gene expression

To provide some insights into the potential molecular mechanisms underlying the TAOassociated variants, we queried whether the five top SNPs and two secondary signals (and their proxies) were 1) tagging potentially deleterious SNPs, 2) located in regulatory elements, and 3) reported to influence the expression of one or more of nearby genes (eQTLs at $P<5x10^{-5}$). We focused on the new TAO risk locus at 16q12 region. Functional annotations for the remaining six loci are presented in the online repository and eQTL data are presentedin Table III (main results) and in Table E3 in the online repository.

The 16q12 TAO-associated variants are located in an intergenic region delimited by two recombination hotspots on each side, near *CYLD* (22 kb downstream). The rs1861760 maps to FOXJ1 and SOX binding sites. This SNP and/or its proxies correlate with the expression of *CYLD* in both blood and human lung tissues and of *NOD2* (*nucleotide-binding oligomerization domain 2*) in blood²⁰ (Table III for main results and Table E3 in the online repository).^{17,20}

322

323 Relationship between TAO-associated variants and age-of-asthma onset

To investigate whether TAO-associated SNPs influence the age-of-asthma onset, we 324 compared, in asthmatics, the distribution of age-of-asthma onset according to the number of 325 326 risk alleles at each of the seven main and secondary TAO-associated SNPs (Figure 3). Asthmatic subjects carrying one or two copies of risk allele at 17q12-q21 SNPs (rs9901146 327 and rs3859192) or at 9p24 rs928413 had a younger age-of-asthma onset than non-carriers 328 (median 6-8 years versus 10 years ($P \le 6x10^{-4}$) and median 6-8 years versus 9 years (P = 0.002), 329 respectively), whereas those having at least one copy of rs1861760 risk-allele at 16q12 had a 330 later age-of-asthma onset than non-carriers (median 10 years versus 8 years; P=0.04). No 331 332 significant difference was found for the other three SNPs. We evidenced that an increased number of risk alleles at these four SNPs was associated with a younger age-of-asthma onset 333 (median 12 years for carrying one risk allele to 4 years for carrying 6-8 risk alleles; $P=10^{-4}$). 334 335 Finally, we detected a strong association between the age-of-asthma onset and the polygenic risk score (from median 10 years in first quintile to 6 years in last quintile; $P=4x10^{-4}$). 336

337

338 Comparison of time-to-asthma onset GWAS results with previous asthma GWASs

To investigate the impact of taking into account the age-of-asthma onset in disease modeling 339 through survival analysis, we explored whether the top TAO SNPs were associated with 340 asthma modeled as a binary trait in the nine cohorts included in the present study (Table E4, 341 see online repository). We also investigated GABRIEL top SNPs in TAO meta-analysis 342 (Table E4).¹⁴ We observed a strong decrease of heterogeneity of SNP effect across studies in 343 TAO analysis ($P_{Het} \ge 0.11$) compared with asthma binary trait analyzed in the same datasets 344 (P_{Het}≥0.004) as well as in all GABRIEL datasets (P_{Het}≥0.0009), especially in 9p24 and 17q12-345 q21 regions. The association signals were always more significant in TAO analysis as 346 compared with the binary trait analysis in the same datasets. This increase of significance 347 level was very high: 100-fold for 2q12 and 16q12, and 10^4 to 10^6 -fold for 9p24 and 17q12-348 q21. In fact, the asthma binary trait analysis only detected two loci (HLA and GSDMA) at the 349 genome-wide significance level out of the seven TAO-associated loci. Conversely, the present 350 351 TAO analysis identified at the genome-wide significance level four of the six main published GABRIEL regions,¹⁴ and event at higher significance for 9p24 and 17q12-q21 regions (100 to 352 10⁴-fold) as compared to GABRIEL significance levels. The two remaining GABRIEL loci 353 not detected by our TAO analysis were those with the weaker effects (OR=1.12 for rs744910 354 in 15q22 and rs2284033 in 22q13) in GABRIEL meta-analysis.¹⁴ 355

Finally, we evaluated whether previously reported susceptibility loci for $asthma^{26}$ were associated with time-to-asthma onset in our meta-analysis (see online repository, Table E5). Among the 21 loci detected in European populations, 12 were replicated at 5% in our TAO meta-analysis with the same direction of effects. Among the nine non-replicated signals, three SNPs (or some proxies) were not available in our data, and the remaining six loci had been reported for specific phenotypes: asthma exacerbation, age-of-asthma onset *per se* in asthmatic children only (quantitative trait) or childhood asthma (binary trait).^{11,27,28}

364 **DISCUSSION**

By taking into account the age-of-asthma onset in asthma association analysis, we identified in this large meta-analysis including both asthmatic and non-asthmatic individuals, adults and children, a new susceptibility locus at 16q12 associated with time-to-asthma onset and confirmed the involvement of six other distinct loci belonging to four regions in asthma pathogenesis (2q12, 6p21, 9p24 and 17q12-q21). Genetic variants at 9p24 and 17q12-q21 were strongly associated with an earlier onset of childhood asthma whereas the 16q12 lead SNP was associated with a risk of later onset asthma.

The most significant 16q12 genetic variant (rs1861760) is located nearby CYLD and NOD2 372 and also maps to a binding site of FOXJ1, a transcription factor associated with allergic 373 rhinitis.²⁹ Genetic variants located in a 130-kb region around rs1861760 were reported 374 associated with immune-related diseases: inflammatory bowel diseases (Crohn's disease) and 375 leprosy.³⁰⁻³⁵ Interestingly, haplotype reconstruction (Figure 4) showed that the TAO 376 377 rs1861760-A risk allele was always associated with SNP alleles that conferred a decreased risk of Crohn's disease (rs17221417-C, rs5743289-C and rs2076756-A located in NOD2 and 378 rs12324931-A located in CYLD)^{30-32,35,36} and of leprosy (rs16948876-G located in intergenic 379 region at 2 kb from rs1861760).³³ Indeed, GWASs revealed common genetic susceptibility 380 loci for asthma and other immune-related disorders, suggesting shared molecular pathways 381 involved in their etiology; however opposite alleles appear to be at risk.³⁷ Interestingly, an 382 opposite effect of rs1861760-A allele is also observed at the gene expression level. Thus, 383 TAO risk-allele at rs1861760 correlated with both the expression of CYLD and NOD2 in 384 blood but with an opposite effect.²⁰ However, this TAO risk allele was only associated with 385 increased CYLD expression in lung tissue.¹⁷ CYLD encodes a deubiquitinating enzyme that 386 regulates diverse physiological processes including immune response and inflammation.³⁸ 387 CYLD mainly acts as a negative regulator of nuclear factor-kB (NFkB1) to protect the host 388

from an over-reactive inflammatory response.³⁸ Conversely NOD2, which plays an important 389 role in the innate immune response to intracellular bacterial lipopolysaccharides (LPS), 390 activates the *NFkB1* pathway.³⁹ NFkB1 is a pleiotropic transcription factor that acts as a key 391 regulator of immune and inflammatory genes, and activation of the NFkB1 pathway has been 392 implicated in airway inflammation and asthma.^{40,41} Moreover, FOXJ1 transcription factor that 393 binds to the genomic region encompassing 16q12 TAO-associated SNP (rs1861760) was 394 described to inhibit NFkB1 activity.⁴² Recently, CYLD has been shown to regulate lung 395 fibrosis in mice by inhibiting transforming growth factor- β -signaling (TGF- β) through a 396 decrease of SMAD3 protein stability.⁴³ Of interest, SMAD3 has been reported to be associated 397 with asthma by previous GWAS.¹⁴ 398

Defining the phenotype is an important consideration because phenotypic heterogeneity can 399 reduce power of GWAS.⁴⁴ In the present analyses, we studied the variability of time-to-400 401 asthma onset in both asthmatics and non-asthmatics based on survival analysis methods. The information used for such analysis was the age-of-onset in asthmatics and the age at last 402 403 examination or death in non-asthmatics. In such a model, unaffected subjects represent censored observations as they are still at risk for disease, being perhaps too young to exhibit 404 the trait. This approach, which allowed combining the age-of-asthma onset together with the 405 disease status (affected/unaffected) led to decrease the genetic heterogeneity across studies 406 and increase the power to detect association signals (upon 10⁶-fold increase as compared to 407 the disease status only analysis. More specifically, increased evidence of association was 408 observed in regions where age-of-asthma onset explained at least in part the genetic 409 heterogeneity such as the 17q12-q21 locus for which a restricted SNP effect to a particular 410 group of age-of-onset (early childhood onset asthma) was demonstrated.¹³ Moreover, this 411 analysis led to the identification of a new locus at 16q12 near CYLD and of an additional 412 signal in the 9p24 region. These results support the hypothesis that a better consideration of 413

the phenotypic heterogeneity of asthma may help to disentangle the genetic heterogeneity ofasthma.

Our study included both children and adult asthmatic subjects. Age-of-disease onset may be 416 417 subject to recall bias, especially among individuals who are the furthest from the time of the first symptoms (e.g. adults who suffered of asthma in childhood) because it is often defined in 418 a retrospective manner. However, high accuracy of the self-reported year of asthma onset by 419 adult subjects has been shown by two independent studies among which the European 420 Community Respiratory Health Survey, which was part of the present study.^{45,46} Erroneous 421 recall of the age-of-asthma onset is unlikely to have significantly affected the results since we 422 observed little genetic heterogeneity across studies (e.g. childhood onset asthma reported by 423 either adults or children). 424

It was suggested that some genetic variants might influence asthma in an age-specific manner. 425 426 Among TAO-associated SNPs, we confirmed the association of 17q12-q21 SNPs with an early age-of-asthma onset^{13,14} and evidenced for the first time that the top 9p24 genetic variant 427 428 near IL33 was also associated with early childhood-onset asthma (median age-of-onset of 6-8 429 years in risk-allele carriers). Indeed, in GABRIEL meta-analysis, 9p24 SNPs were more strongly associated with early-onset asthma (before 16 years) than late-onset asthma (after 16 430 years), but this difference was not significant.¹⁴ Conversely, genetic variants at the new 431 susceptibility locus, 16q12, conferred a risk of later onset asthma (median age-of-onset of 10 432 years in risk-allele carriers). Moreover, we evidenced that a high burden of disease risk alleles 433 at these loci is associated with earlier age-of-asthma onset (4 years versus 9-12 years). This 434 difference in asthma onset may reflect difference in patterns of onset of disease.⁴⁷ Indeed, we 435 evidenced in the GABRIELA study that persistent early wheezing subjects carried more risk 436 437 alleles than transient early wheezing subjects, and we confirmed previous association between persistent early wheezing and 9p24 and 17q12-q21 loci (data not shown). The 17q12-q21 438

genetic variants were reported associated with persistent childhood wheeze phenotype
whereas 9p24 were mostly associated with intermediate onset wheeze but also with persistent
early wheeze.^{48,49} Moreover, 17q12-q21 SNPs were associated with fractional exhaled Nitric
Oxide (FeNO) levels in children but not in adults, childhood severe asthma and allergic
rhinitis, and 9p24 SNPs were associated with childhood severe asthma, asthma-plus-rhinitis,
atopic asthma, allergy and eosinophil counts.⁵⁰⁻⁵⁵

In summary, we identified five regions harboring seven distinct signals that were associated with time-to-asthma onset, among which the 16q12 region that is reported for the first time. Several lines of evidence suggest that *CYLD* and *NOD2* located in that region are strong candidate genes for asthma. This study demonstrates that incorporating the variability of age-

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657

Table I. Top SNPs in main loci associated with time-to-asthma onset at genome-wide ($P \le 5x10^{-8}$) and suggestive significance levels ($5x10^{-6}$) $^{8} < P < 10^{-6}$)

						Time to asthma onset - N=13,886			
Chromosome	Marker	Position [*]	Nearest gene(s) (kb distance)	Effect/ Ref Alleles [†]	Effect Freq	Hazard Ratio [95% CI]	\mathbf{P}^{\ddagger}	P-Het ^{**}	
Loci with geno	ome-wide signif	ïcance - P≤5x1)-8						
2q12	rs10208293	102966310	IL1RL1 ^{***}	G/A	0.73	1.14 [1.08-1.19]	3.1x10 ⁻⁸	0.26	
6p21	rs9272346	32604372	HLA-DQA1 (0.8)	A/G	0.59	1.13 [1.08-1.17]	1.6x10 ⁻⁸	0.12	
9p24	rs928413	6213387	IL33 (2)	G/A	0.25	1.19 [1.13-1.25]	6.5x10 ⁻¹⁶	0.15	
16q12	rs1861760	50857693	CYLD (22)	A/C	0.04	1.28 [1.17-1.40]	4.2x10 ⁻⁸	0.11	
17q12-q21	rs9901146	38043343	ZPBP2 (9)	G/A	0.51	1.18 [1.13-1.22]	1.9x10 ⁻¹⁶	0.17	
			GSDMB (17)						
Suggestive loci	i - 5x10 ⁻⁸ <p<10< td=""><td>-6</td><td></td><td></td><td></td><td></td><td></td><td></td></p<10<>	-6							
2q11-q12	rs12468899	102426140	MAP4K4 ^{***}	G/A	0.69	1.12 [1.09-1.16]	1.7×10^{-7}	0.89	
15q22	rs11071559	61069988	RORA ^{***}	C/T	0.85	1.16 [1.10-1.24]	8.3x10 ⁻⁷	0.96	
16p12-p11	rs1805013	27373980	IL4R ^{***}	T/C	0.05	1.22 [1.13-1.32]	8.0x10 ⁻⁷	0.37	

^{*}Position in base pairs (bp) – build 37.3 NCBI

- [†]For the calculation of the hazard ratios, effect alleles were designated as risk alleles. Effect Freq denotes frequency of the effect allele, CI
- 663 confidence interval, and Ref reference allele.
- ⁶⁶⁴ [‡]P-value obtained from single-SNP Cox model for time-to-asthma onset adjusted for sex and principal components (fixed-effect model when
- there was no significant evidence of heterogeneity or random-effect model otherwise)
- 666 **P-Het reflects P-value of Cochran's Q statistic across studies
- 667 **** SNP is located within reported gene
- 668

Table II. Secondary signals associated with time-to-asthma onset after stepwise conditional analysis in 9p24 and 17q12-q21 regions. This table
contains, for these two regions, the top time-to-asthma onset SNP in bold (rs928413 and rs9901146 respectively) and the most significant SNP in
the conditional analysis, after fitting the lead SNP in the region.

						Single SNP analysis			Fitted SNP(s)			
Chr	Marker	Nearest gene (kb distance)	Position*	Effect/Ref Alleles [†]	Effect Freq	Hazard Ratio	\mathbf{P}^{\ddagger}	P-Het ^{**}	Hazard Ratio	\mathbf{P}^{\ddagger}	P-Het ^{**}	
						[95% CI]			[95% CI]			
9p24 r	egion								rs	928413		
9	rs413382	IL33 (73)	6142948	A/C	0.80	1.15	5.9x10 ⁻⁸	0.84	1.13	9.7x10 ⁻⁶	0.80	
						[1.08-1.22]			[1.06-1.20]			
9	rs928413	IL33 (2)	6213387	G/A	0.25	1.19	6.5x10 ⁻¹⁶	0.15	-	-	-	
						[1.13-1.25]						
17q12-q21 region									rs9	901146		
17	rs9901146	ZPBP2 (9)	38043343	G/A	0.51	1.18	$1.9 \mathrm{x10}^{-16}$	0.17	-	-	-	
						[1.13-1.22]						
17	rs3859192	GSDMA ^{***}	38128648	T/C	0.48	1.16	1.5×10^{-13}	0.90	1.11	4.0×10^{-6}	0.74	
						[1.12-1.21]			[1.06-1.15]			

⁶⁷² *Position: Position in base pairs (bp) - build 37.3 NCBI

- ⁶⁷³ [†]For the calculation of the hazard ratio, effect alleles were designated as risk alleles. Effect Freq denotes frequency of the effect allele, CI
- 674 confidence interval, and Ref reference allele
- ⁴P-values are obtained from Cox model of time-to-asthma onset adjusted for sex and principal components
- ⁶⁷⁶ ^{**}P-Het reflects P-value of Cochran's Q statistic across studies
- 677 **** SNP is located within reported gene
- 678

Table III. Main cis-eQTLs results for the top SNPs in genome-wide associated regions from the meta-analysis of time-to-asthma onset. We

680 focused on eQTLs measured in blood, lymphoblastoïd cell lines (LCLs) and lung tissue.

Locus	\mathbf{SNP}^*	Alleles	Gene(s)	Range of P-values	Tissue	Source ^{††}
	(LD with top SNP)	(Ref/Effect)				
2q12	rs10208293	G/A	IL18RAP, IL18R1	2.5x10 ⁻¹³ -9.8x10 ⁻¹⁹⁸	Blood, LCLs	Blood eQTLs, eQTL Browser
6p21	rs9272346	G/A	HLA-DQA1/DQA2/DQAS1/ DQB1/DQB2,HLA-DRA/DRB1/ DRB5/DRB6, TAP2	1.3x10 ⁻⁶ -2.1x10 ⁻¹²¹	LCLs, Lung, Blood	eQTL_Chicago,GTEx, Blood eQTLs
16q12	rs1861760	C/A	NOD2	3.6x10 ⁻¹¹	Blood	Blood eQTLs
	rs5743266 [†]		CYLD, NOD2	$5.0 \times 10^{-9} - 3.2 \times 10^{-120}$	Blood	Blood eQTLs
	(D'=1, r ² =0.02)					
	rs7205760 [†]		CYLD, NOD2	2.8x10 ⁻⁶ -4.0x10 ⁻¹⁵	Lung, Blood	Lung eQTLs, Blood eQTLs
	(D'=1, r ² =0.005)					
17q12-q21	rs9901146	A/G	GSDMB, ORMDL3	3.8x10 ⁻⁶ -9.8x10 ⁻¹⁹⁸	Blood, LCLs	Blood eQTLs, GTEx, eQTL Browser, eQTL_Chicago
	rs3859192	C/T	GSDMA, GSDMB, ORMDL3	$1.1 \times 10^{-7} - 2.5 \times 10^{-12}$	Lung, LCLs	GTEx, eQTL Browser

682 indicated in bold.

- [†]Haplotype reconstruction was done using Haploview; the effect allele of the top SNP (A-rs1861760) is always transmitted with the effect allele
 of its proxy (G-rs5743266 and G-rs7205760).
- ^{††}Interrogated databases: *eQTL Browser* (LCLs of British subjects with asthma or eczema),¹⁸ *Blood eQTL Browser* (non-transformed peripheral
- blood sample),²⁰ Lung eQTLs (lung),¹⁷GTEx eQTL Browser v4 (several tissues among which blood and lung),²³ and eQTL Chicago Browser (LCLs).^{19,21,22}

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Figure 1. Manhattan plot showing the association P-values of the genome-wide association results for time-to-asthma onset from the meta-analysis. The $-\log_{10}$ of the P-value for each of 2,400,368 SNPs (y-axis) is plotted against the genomic position (x-axis). The solid red line indicates the genome-wide significance threshold of P=5x10⁻⁸.

Figure 2. Regional association plot of the 16q12 region using Locuzoom software.⁵⁶ SNPs are plotted with their P-values ($-\log_{10}$ values, left y-axis) as a function of genomic position (xaxis). Estimated recombination rates (right y-axis) taken from 1000G (EUR) are plotted to reflect the local LD structure around the top associated SNP (purple circle) and correlated proxies (according to a blue to red scale from r²=0 to 1).

Figure 3. Relationship between time-to-asthma onset associated SNPs and age-of-asthma
onset. Part A) Association between age-of-asthma onset and genotypes at individual locus;
Part B) Median of age-of-asthma onset as a function of the number of individual's risk allele
burden; Part C) Median of age-of-asthma onset by quintile of genetic risk score.

Figure 4. Map of 16q12 region (build 37.3 position: 50,723,355-50,860,722) and haplotype reconstruction for SNPs found associated with inflammatory bowel disease (among which Crohn's disease, blue), leprosy (green) or asthma (red) or with the expression of *CYLD* or *NOD2* (black). Linkage disequilibrium plot was obtained using Hapmap2 CEU reference sample from Haploview⁵⁷ (values and colors reflect r² and D' respectively). The 16q12 top SNP (rs1861760) associated with time-to-asthma onset is indicated in bold.

Figure 1.



Figure 2.









