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Anthropometry, breast cancer and menopausal status: use of repeated measurements over 10 years of follow-up – Results of the French E3N women's cohort study

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Key words: BMI; weight; breast cancer; methodological issue; HRT use; cohort study

The association between weight, BMI and breast cancer was analyzed on 94,805 women of the E3N cohort according to their menopausal status. Seven hundred eighty-six incident invasive premenopausal breast cancers and 1,522 incident invasive postmenopausal breast cancers occurred during a mean follow-up of 9.7 years. Weight and BMI were updated every 24 months and considered as time-dependent variables. Data were analyzed using multivariate Cox proportional hazards models. Trend RRs of premenopausal breast cancer were 0.97 (0.92–1.01) for a 5 kg increase in weight and 0.96 (0.91–1.01) for a 2 kg/m² increase in BMI, adjusted for other known risk factors. Opposite trend RRs were found after menopause: 1.05 (1.02–1.08) for weight and 1.06 (1.02–1.09) for BMI, respectively, for similar increases. Women with a BMI of over 30 kg/m² had a RR of premenopausal breast cancer of 0.66 (0.40–1.10) compared to those with a BMI of between 18.5 and 25 kg/m². Postmenopausal women with a BMI of over 30 kg/m² had a RR of breast cancer of 1.23 (1.00–1.59). The increase in risk of postmenopausal breast cancer with increased weight or BMI was similar whatever the HRT used, although the point estimates were higher in HRT users. We strongly recommend to use anthropometric measurements updated during follow-up to assess the effect of weight, BMI on breast cancer risk.

Literature reviews^{1–3} and meta-analysis⁴ showed evidence of a positive relation between overweight or obesity and postmenopausal breast cancer and a possible inverse relation with premenopausal breast cancer.

Hypothesizing that overweight affects hormonal metabolism throughout life, and given the sharp increase in obesity in most Western countries, our aim was to examine weight and body mass index (BMI) in relation to pre- and post-menopausal breast cancer occurrence. Moreover, as weight and BMI might vary through life, especially at menopause, we considered it essential to analyze them as time-dependent variables allowing adequate control of their variations. Only a few cohort studies^{2,5,6} have analyzed weight and BMI over time and they limited their investigation of variations in weight or BMI to a comparison between measurements made at the beginning and end of follow-up. In order to avoid errors due to risk exposure only measured at enrollment in the cohort study, *i.e.*, possibly far from the event of interest, we modeled breast cancer risk using anthropometric measurements regularly updated.

E3N (Etude Epidémiologique de femmes de la Mutuelle Générale de l'Education Nationale), a prospective cohort study on French women, offered the opportunity to study the evolution of weight and BMI in relation to pre- and post-menopausal breast cancer occurrence over a 10-year follow-up period, during which weight was recorded every 2 years.

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Material and methods

E3N's main objective is to investigate risk factors for cancer. The design of the study has been described elsewhere.⁷ Briefly, the cohort consists of 98,997 women living in France, who were insured with a national health insurance scheme primarily covering teachers. They were aged 40–65 years at enrollment (between June 1990 and November 1991) after replying to a baseline questionnaire. The E3N cohort is the French part of the European Prospective Investigation on Cancer (EPIC8).

Since baseline, participants were followed by self-administered questionnaires sent out approximately every 24 months. We analyzed the variables on weight as self-reported in each of the first 5 questionnaires and on height as recorded at baseline.

Menopause was recorded in each follow-up questionnaire. To ensure that the constructed menopause variables were as accurate as possible, the whole set of answers on date and type of menopause (natural or the result of bilateral oophorectomy, chemotherapy, radiotherapy or other treatment), date of last menstruation, date of start of menopausal symptoms and date of hysterectomy, if appropriate, were taken into account. Postmenopause was defined as the cessation of periods for natural reasons or due to radiation, chemotherapy or surgery (total oophorectomy). The menopausal status of 24,910 women changed during follow-up. Women contributed to the premenopausal group for the period between enrollment and onset of menopause and then to the postmenopausal group until the end of follow-up. Women with undefined menopausal status (for instance, because of continuous use of hormonal treatments or hysterectomy with no additional information on oophorectomy) were not considered in the analysis ($n = 1790$) nor were those who had never menstruated ($n = 4$). Women who had reported a cancer other than a basal cell carcinoma at enrollment were excluded ($n = 2,398$).

All questionnaires asked participants whether breast cancer had been diagnosed, requesting the addresses of their physicians and permission to contact them. Information on nonrespondents was obtained from the MGEN file on the reimbursement of hospital fees, enabling additional breast cancer cases to be found.

Follow-up time is considered between return of the baseline questionnaire and June 28, 2000, the date at which the 6th questionnaire was sent out. Mean follow-up is 9.7 years (Std = 1.2 years). The present analysis is based on the follow-up of a total sample of 94,805 women who between them reported 2,308 breast cancers, consisting of 786 premenopausal and 1,522 postmenopausal breast cancers.

Data were analyzed using Cox proportional hazards models with age as the time scale. The adjustment variables taken into account were age at menarche (<12, 12, 13 and ≥ 14), age at first birth (<23, 23–25, 26–29 and ≥ 30), parity (0, 1–3 and ≥ 4), history of breast cancer in first degree relatives (yes/no), history of benign breast disease (yes/no), alcohol consumption (drink(s) per week, 0, <1 and ≥ 1), number of years at school (0, 1–5, 6–9, 10–13, 14–15 and ≥ 16) and ever married (yes/no). Additional adjustments were performed on quartiles of physical activity and height. Use of HRT was tested as a potential effect modifier. Weight and BMI were updated and introduced in the models as time-dependent variables. Women with missing values on weight or BMI were excluded from calculation during the time for which these values were missing. Weight and BMI were considered as continuous variables and categorized in quartiles according to their distribution in the replies to the first questionnaire. We also considered WHO cut-off points for the BMI.⁹

All analyses were performed with SAS[®] Software.¹⁰

Results

The anthropometric characteristics of cases and noncases, based on the replies to each questionnaire, were compared (Table I). Marginal differences were observed between cases and noncases. Among premenopausal women, cases were slightly leaner than noncases, but mean differences on weight and BMI did not exceed 1.6 kg and 0.7 kg/m² respectively. In contrast, among postmenopausal women, cases were slightly heavier than noncases with maximum mean differences lower than 1.8 kg and 0.5 kg/m².

Table 1: Anthropometric characteristics of breast cancer cases¹ and non-cases in the E3N cohort by menopausal status* at the time of each questionnaire.

| Variable | First questionnaire | | Second questionnaire | | Third questionnaire | | Fourth questionnaire | | Fifth questionnaire | |
|------------------------------|---------------------|------------|----------------------|-------------|---------------------|-------------|----------------------|------------|---------------------|--------------|
| | N | Mean (Std) | N | Mean (Std) | N | Mean (Std) | N | Mean (Std) | N | Mean (Std) |
| Premenopausal women | | | | | | | | | | |
| Weight (kg) | | | | | | | | | | |
| Cases | 181 | 59.3 (9.9) | 153 | 58.6 (8.9) | 120 | 59.1 (8.1) | 104 | 59.1 (9.6) | 134 | 59.2 (8.1)* |
| Noncases | 48,963 | 58.2 (8.7) | 36,485 | 58.8 (9.0) | 26,357 | 58.9 (8.7) | 19,896 | 59.8 (9.5) | 20,043 | 60.8 (10.1) |
| BMI (kg/m ²) | | | | | | | | | | |
| Cases | 181 | 22.3 (3.5) | 153 | 22.2 (3.1) | 120 | 22.4 (2.6) | 104 | 22.4 (3.1) | 134 | 22.4 (2.7)†* |
| Noncases | 48,963 | 22.1 (3.0) | 36,485 | 22.3 (3.1) | 26,357 | 22.4 (3.0) | 19,896 | 22.6 (3.3) | 20,043 | 23.1 (3.6) |
| Post-menopausal women | | | | | | | | | | |
| Weight (kg) | | | | | | | | | | |
| Cases | 156 | 60.3 (9.8) | 206 | 61.8(10.2)† | 150 | 62.1(11.1)† | 239 | 61.8 (9.7) | 560 | 62.0 (9.7) |
| Noncases | 41,271 | 60.0 (9.4) | 41,469 | 60.4 (9.7) | 38,117 | 60.3 (9.2) | 41,215 | 61.0 (9.8) | 61,847 | 61.8 (10.3) |
| BMI (kg/m ²) | | | | | | | | | | |
| Cases | 156 | 23.3 (3.3) | 206 | 23.6 (3.6) | 150 | 23.6 (3.9) | 239 | 23.7 (3.6) | 560 | 23.7 (3.5) |
| Noncases | 41,271 | 23.1 (3.6) | 41,469 | 23.2 (3.4) | 38,117 | 23.1 (3.3) | 41,215 | 23.4 (3.5) | 61,847 | 23.7 (3.7) |

¹ Incident cases occurring between two consecutive questionnaires

* Menopausal status at enrolment in the study

†: p-value < 0.05.

Variations in BMI for pre- and post-menopausal women during follow-up were studied. A positive trend was seen in BMI gain since enrollment in the study in 1990 both in the premenopausal (Fig. 1) and postmenopausal (Fig. 2) groups, whatever the age category. After menopause, the increase in BMI was apparent, though it was more pronounced in the 2 younger birth cohorts.

Figure 1: BMI variations during follow-up, according to age categories at inclusion, for premenopausal women of the E3N cohort.

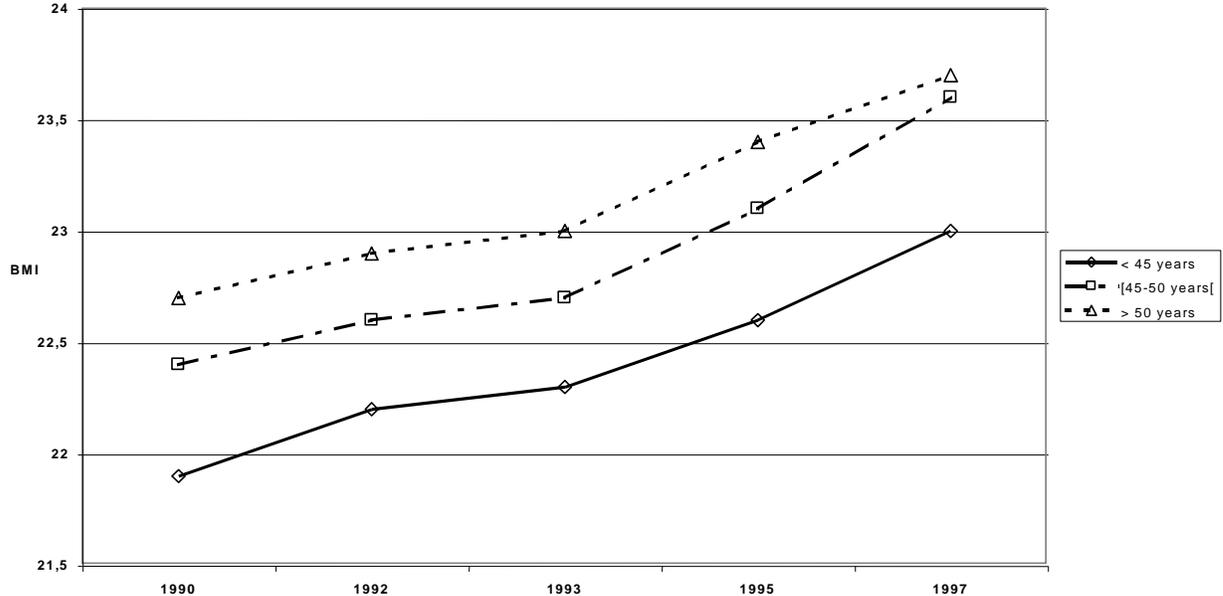
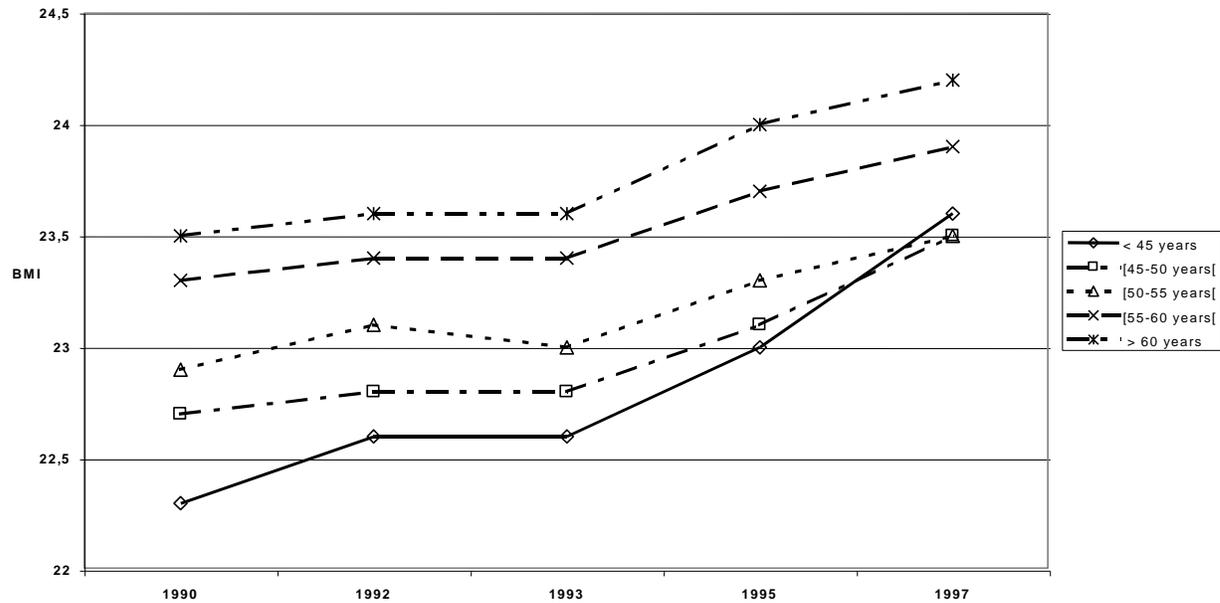


Figure 2: BMI variations during follow-up, according to age categories at inclusion, for postmenopausal women of the E3N cohort.



Relative risks of breast cancer associated with weight and BMI as time-dependent variables are shown in Table II. A decrease in risk of premenopausal breast cancer with increasing weight and BMI was observed, though no trend test reached significance (trend RR = 0.97 (0.92–1.01) for a 5 kg increase, trend RR = 0.96 (0.91–1.01) for a 2 kg/m² increase). Relative risks in the fourth quartile of weight and BMI, as compared to the first, were 0.75 (0.61–0.93) and 0.78 (0.64–0.94), respectively. The relative risk of premenopausal breast cancer reached 0.66 (0.40–1.10) for women with a BMI of over 30 kg/m² compared to women with a BMI between 18.5 and 25 kg/m².

Significant positive trends of increasing risk of postmenopausal breast cancer with both increasing weight and BMI were observed (trend RR = 1.05 (1.02–1.08) for weight and trend RR = 1.06 (1.02–1.09) for BMI, respectively). The RR associated with weight and BMI for the fourth quartile, as compared to the first, were 1.06 (0.91–1.21) and 1.06 (0.93–1.21), respectively. Because the cut-offs were defined according to the baseline distribution, the number of subjects in the upper quartile increased during follow-up. We therefore divided it into 2 categories. Higher RRs were observed for the upper categories of weight [over 68 kg: RR = 1.10 (0.93–1.29)] and BMI [over 26.2 kg/m² RR = 1.15 (1.00–1.34) as compared to the first quartiles]. Women with a BMI of under 18.5 kg/m² had a lower risk, of 0.72 (0.51–1.00), while a RR of 1.23 (1.00–1.59) was seen in obese women (BMI over 30 kg/m²) as compared to women with a BMI of between 18.5 and 25 kg/m².

Interactions between weight or BMI and HRT use were studied in relation to postmenopausal breast cancer occurrence because of a possible differential effect of HRT use in the relation between weight or BMI and breast cancer. HRT use was recorded in each follow-up questionnaire and we considered it as a time-dependent variable in our models. Significant positive trend RRs were observed (Table III) in both subgroups (HRT users and HRT neverusers). Trend RRs related to BMI were equal to 1.06 (1.01–1.12) for HRT users and 1.05 (1.01–1.10) for HRT never users. There was however some indication of a possible effect modification by HRT use, with a weak difference in risk estimates in the highest quartiles of BMI.

Table II: Relative risks of pre- and post-menopausal breast cancer according to weight and BMI self-reported every two years. E3N cohort (1990-2000).

| Variable | Cases | Total person-years | Multivariate relative risk | Trend RRs ¹ |
|------------------------------|-------|--------------------|----------------------------|------------------------|
| Premenopausal women | | | | |
| Weight (kg) | | | | |
| Q1 (≤ 52) | 153 | 65,0861 | 1.00 (reference) | 0.97 (0.92–1.01) |
| Q2 (52–57) | 173 | 73,041 | 0.86 (0.70–1.04) | |
| Q3 (57–63) | 202 | 72,236 | 0.96 (0.79–1.17) | |
| Q4 (> 63) | 164 | 73,180 | 0.75 (0.61–0.93) | |
| BMI (kg/m ²) | | | | |
| Q1 (≤ 20.2) | 162 | 64,986 | 1.00 (reference) | 0.96 (0.91–1.01) |
| Q2 (20.2–21.6) | 161 | 65,782 | 0.86 (0.71–1.05) | |
| Q3 (21.6–23.4) | 185 | 71,747 | 0.90 (0.74–1.09) | |
| Q4 (>23.4) | 184 | 81,028 | 0.78 (0.64–0.94) | |
| BMI (kg/m ²) | | | | |
| < 18.5 | 33 | 12,697 | .06 (0.74–1.50) | 2 |
| [18.5–25[| 553 | 225,058 | 1.00 (reference) | |
| [25–30[| 91 | 36,968 | 0.93 (0.74–1.16) | |
| ≥ 30 | 15 | 8,820 | 0.66 (0.40–1.10) | |
| Post-menopausal Women | | | | |
| Weight (kg) | | | | |
| Q1 (≤ 53) | 254 | 98,846 | 1.00 (reference) | 1.05 (1.02-1.08) |
| Q2 (53-58) | 297 | 106,508 | 0.95 (0.82-1.10) | |
| Q3 (58-64) | 331 | 109,358 | 1.02 (0.88-1.18) | |
| Q4 (> 64) | 429 | 140,395 | 1.06 (0.91-1.21) | |
| Q4 ₁ (64-68) | 158 | 54,263 | 0.99 (0.82-1.19) | 1.10 (0.93-1.29) |
| Q4 ₂ (>68) | 271 | 86,131 | 1.10 (0.93-1.29) | |
| BMI (kg/m ²) | | | | |
| Q1 (≤ 20.6) | 239 | 87,955 | 1.00 (reference) | 1.06 (1.02-1.09) |
| Q2 (20.6-22.2) | 272 | 96,899 | 0.91 (0.79-1.07) | |
| Q3 (22.2-24.2) | 331 | 117,463 | 0.95 (0.81-1.08) | |
| Q4 (>24.2) | 469 | 152,791 | 1.06 (0.93-1.21) | |
| Q4 ₁ (24.2-26.2) | 202 | 70,906 | 0.97 (0.81-1.14) | 1.15 (1.00-1.34) |
| Q4 ₂ (>26.2) | 267 | 81,884 | 1.15 (1.00-1.34) | |
| BMI (kg/m ²) | | | | |
| < 18.5 | 33 | 15,135 | 0.72 (0.51-1.00) | 2 |
| [18.5-25[| 916 | 320,722 | 1.00 (reference) | |
| [25-30[| 285 | 95,580 | 1.05 (0.92-1.20) | |
| ≥ 30 | 77 | 10,943 | 1.23 (1.00-1.59) | |

¹Trend RRs are calculated on a 5 kg increase in weight and a 2 kg/m² increase in BMI. ²The test is identical for both variables on BMI.

Table III: Relative risks of post-menopausal breast cancer according to weight and BMI self-reported every two years in the E3N cohort, by HRT use¹. E3N cohort (1990-2000).

| Variable | HRT users | | | HRT never users | | |
|-------------------------------|-------------------------------|-------------------------------|------------------------|-------------------------------|-------------------------------|------------------------|
| | Total (cases) person-years | Multivariate relative risk | Trend RRs ² | Total (cases) person-years | Multivariate relative risk | Trend RRs ² |
| Weight (Kg) | | | | | | |
| Q1 (≤ 53) | 49,158 (135) | 1.00 (reference) | 1.06 (1.01-1.11) | 49,687 (119) | 1.00 (reference) | 1.04 (1.00-1.09) |
| Q2 (53-58) | 54,540 (157) | 1.00 (0.82-1.25) | | 51,967 (140) | 0.90 (0.73-1.11) | |
| Q3 (58-64) | 54,591 (185) | 1.17 (0.95-1.44) | | 54,766 (146) | 0.89 (0.72-1.10) | |
| Q4 > 64 | 58,804 (203) | 1.19 (0.97-1.48) | | 81,590 (226) | 0.94 (0.78-1.15) | |
| BMI (Kg/m²) | | | | | | |
| Q1 (≤ 20.6) | 44,463 (122) | 1.00 (reference) | 1.06 (1.01-1.12) | 43,491 (117) | 1.00 (reference) | 1.05 (1.01-1.10) |
| Q2 (20.6-22.2) | 50,854 (149) | 1.02 (0.82-1.26) | | 46,044 (123) | 0.85 (0.68-1.04) | |
| Q3 (22.2-24.2) | 58,542 (179) | 1.06 (0.86-1.30) | | 58,920 (152) | 0.83 (0.67-1.01) | |
| Q4 (>24.2) | 63,234 (230) | 1.26 (1.04-1.54) | | 89,556 (239) | 0.90 (0.75-1.09) | |
| BMI (Kg/m²) | | | | | | |
| < 18.5 | 6,894 (12) | 0.56 (0.32-0.99) | 3 | 8,239 (21) | 0.88 (0.57-1.35) | 3 |
| [18.5-25[| 162,573 (500) | 1.00 (reference) | | 158,148 (416) | 1.00 (reference) | |
| [25-30[| 40,327 (140) | 1.12 (0.93-1.35) | | 55,252 (145) | 1.00 (0.82-1.20) | |
| ≥ 30 | 3,323 (28) | 1.32 (0.90-1.95) | | 16,372 (49) | 1.17 (0.89-1.57) | |

¹Considered as a time-dependent variable. Women are in the nonusers group until first use. ²Trend RRs are calculated on a 5 kg increase in weight and a 2 kg/m² increase in BMI. ³The test is identical for both variables on BMI.

Discussion

The results of our cohort study on women aged between 40 and 65 years showed that overweight had opposite effects on breast cancer risk according to menopausal status over a 10-year follow-up period, taking into account weight change prospectively. We found that the risk of premenopausal breast cancer decreased significantly with increasing weight or BMI, that the risk of postmenopausal breast cancer increased with increasing weight and BMI and lastly that the latter increase was similar in HRT users and never users.

Epidemiological concerns

From the studies published so far, evidence of any negative relationship between premenopausal breast cancer and overweight is unclear.^{2,4,5,11,12} In particular, most results from cohort studies have been nonsignificant.^{2,4} Significance was reached in the Pooling Project on Diet and Cancer, which combined 7 cohort studies,⁴ [RR = 0.90 (0.83–0.97) for a 10 kg increase and RR = 0.89 (0.81–0.97) for a 4 kg/m² increase, respectively]. Our results are lower and consistent with those found in a recent case-control study,¹¹ with RRs of 0.81 (0.55–1.19) and 0.69 (0.47–1.02) for the last quartiles of weight and BMI, respectively, as compared to the first. Other recent case-control studies have found only a weak nonsignificant relation between weight, BMI and premenopausal breast cancer risk¹² or no significant relation at all between weight, BMI and premenopausal breast cancer.⁵

Our results concerning postmenopausal breast cancer risk showed positive trends in risk with weight and BMI, with RRs slightly lower than in previous studies.² In the Pooling Project on Diet and Cancer,⁴ the trend RRs found (trend RR = 1.06 (1.03–1.10) for a 10 kg increase and trend (RR = 1.07 (1.02–1.11) for a 4 kg/m² increase, respectively) were lower than ours. More recently, a case-control study¹¹ found no significant relation between postmenopausal breast cancer risk and weight and BMI. Other recent studies have found high risks with overweight: in a Swedish cohort study, Lahmann *et al.*⁶ found significant increases in risk of postmenopausal breast cancer, with RRs of 1.53 (0.97–2.41) and 1.54 (1.01–2.35) for the highest quintile of weight and BMI compared to the lowest. Case-control studies on postmenopausal Asian women^{5,13,14} have found an approximate doubling in risk for the highest categories of weight and BMI.

We did not confirm the previous findings concerning an effect of anthropometry on postmenopausal breast cancer limited to HRT never users.^{2,4–6,15,16} Abdominal obesity causes an increase of insulin-like growth factor I (IGF-I) activity and a decrease in sex hormone binding globulin (SHBG) level, which are both risk factors of breast cancer. The use of oral estrogen replacement therapy may balance the increase in breast cancer risk due to obesity by contrasting insulin resistance and inducing an increase in SHBG levels and a decrease in circulating IGF-I activity,^{17,18} thus explaining the stronger association between the use of oral unopposed estrogen replacement therapy and breast cancer among lean women as compared to obese ones.¹⁹ In our cohort, 78% of HRT users used a combination of estrogens and progestogens as their main HRT (more than 70% of total HRT use) and the type of progestogens used (androgenic, whether derived from 19-Nor-Testosterone or not, *i.e.*, pregnane or norpregnane derivatives) may have a different influence on insulin levels, IGF1 activity and SHBG levels, accounting for our results.^{17,18}

French women in our cohort are lean as compared to participants of cohort studies in other countries,²⁰ and the HRT users among them had a significantly lower weight and BMI than never users ($p < 10^{-4}$) at the time of reply to each questionnaire. This results in an over-representation of lean women among HRT users and an over-representation of overweight women among HRT nonusers.

Many studies, as reviewed by Friedenreich,² have investigated the potential effect of overweight in breast carcinogenesis. They have found with both pre- and post-menopausal women that obesity is associated with increased plasma concentrations of testosterone and decreased concentrations of SHBG, and as a result of free estrogens. All of these are risk factors for breast cancer and account for the increase found in postmenopausal women. Potischman²¹ observed that serum total estradiol levels decrease with premenopausal obesity, whereas they increase with postmenopausal obesity, which could explain the inverse relation between obesity and breast cancer risk according to menopausal status. In naturally estrogenized premenopausal women, anovulation due to obesity would result in a decrease in estrogen and progesterone production.^{22–24}

Methodological concerns

One major problem for prospective studies occurs when anthropometric measurements made at baseline are used to predict a risk of disease occurring several years later.^{2,25} It may create important discrepancies in results between different studies because in women the rate of increase of weight changes over time, especially with age and menopause. E3N is the first cohort study to consider weight and BMI as time-dependent variables with regularly updated data. In our cohort, 24,910 women changed their menopausal status between their enrollment and 2000, during which period 350 cases of postmenopausal breast cancer occurred. For these women, any estimate of their RR of postmenopausal breast cancer, using their weight and BMI recorded before menopause, would have led to misclassification.

Moreover, some studies have found that the risk relation between postmenopausal breast cancer occurrence and anthropometry is higher for older women.^{4,13,26,27} Consequently, anthropometric measurements made at baseline would underestimate the RRs, especially if the disease is distant from baseline. Updating our data regularly reduces the time lag between the latest measurements and breast cancer occurrence.

As recommended by Korn,²⁸ we used age as the time scale, which means that women were included with increasing ages in the log-likelihood of our Cox model. In such a model, relative risks are calculated within women of equivalent ages. Moreover if age is used as the time scale, no hypothesis on the log-linearity of age is required. On the contrary, in models using time-on-study as the time scale, with adjustment for age, such a hypothesis on the log-linearity of age may be false if, as previously noted, the risk increases with age. On the other hand, the use of age as the time scale, with women of the same age taken into account in the log-likelihood of the model, may introduce an important bias in the evaluation of relative risks in studies using only anthropometric measurements at baseline, since measurements correspond to different ages in these studies. We were able to avoid these 2 major pitfalls by using models for which the anthropometric measurements were updated during the follow-up.

When considering exclusively weight and BMI at enrollment, RRs of premenopausal breast cancer were closer to unity than those obtained with measurements as time-dependent variables (data not shown). Accurate RRs calculation that would take account of the variation of anthropometric characteristics over time may thus reveal a beneficial effect of overweight in the premenopause more important than usually considered.

RRs of postmenopausal breast cancer associated with measurements at enrollment were of a higher magnitude than those with measurements as time-dependent variables (data not shown). The relatively young age of the E3N postmenopausal participants may in part be responsible for our lower estimates as compared to previous published studies, if the effect of overweight on breast cancer risk increases with age. A longer follow-up with further assessments of weight and BMI will also improve our understanding of the effect of anthropometry according to HRT use. Part of the previous published results on postmenopausal breast cancer risk may be over-estimated due to less accurate models.

Limitations of our study

The E3N population is more health-conscious than the general population, with an under-representation of obese women, as compared for example to other European EPIC populations.²⁰ As in most large cohort studies, analyses are based on self-reported anthropometric measurements. Many studies have shown that obese subjects tended to under-estimate their real weight.²⁹⁻³⁷ Nevertheless, recent studies have shown good correlations between self-reported anthropometric measurements and measurements made by technicians.³⁸⁻⁴¹ We conducted a validation study⁴² on 152 women in the E3N cohort, who self-reported their anthropometric characteristics on a questionnaire before the same characteristics were measured by technicians during a short interview. We found no significant differences between reported and measured values, and correlations between reported and measured characteristics reached 0.89 for height, 0.94 for weight and 0.92 for BMI. We also found that women weighing over 65 kg had a pessimistic perception of their body shape less often than leaner women. However, as a result of the prospective design of our study, this would only bias our results towards unity and would in no way produce significant results.

Conclusion

Our study gives evidence that overweight decreases premenopausal breast cancer risk and increases postmenopausal breast cancer risk. The relationship might however be more important in premenopausal women

and less important in postmenopausal women than usually accepted from the literature. Reasons for this discrepancy may come from the capability of our study to take into account variations of anthropometry through life.

When considering only baseline measurements, our results were closer to the literature. Whether this is the consequence of an inadequate modeling of weight variations or whether it is the effect of an increasing effect of overweight with age is debatable.

Our cohort is still ongoing, and a longer follow-up may help to solve this issue. Additional results from cohort studies with repeated measurements of weight are needed. We recommend investigators to analyze these data with careful assessment of variations in weight. A less detrimental effect of overweight on postmenopausal breast cancer risk than usually admitted, and a more marked protective effect before the menopause should not however mask its general unfavorable effects on health.

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